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The extended Weibull distribution with its properties, estimation and modeling skewed data



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ABSTRACT

The present article introduces a new distribution called the odd log-logistic Lindley-Weibull (OLLLW) distribution that provides greater flexibility in modeling data in applied areas such as medicine and engineering. The OLLLW model provides left-skewed, symmetric, right-skewed, and reversed-J shaped densities. Its hazard function can be bathtub, unimodal, increasing or decreasing. The OLLLW density was expressed as a linear mixture of Weibull densities. Some distributional properties of the introduced model are derived. Its parameters are estimated using five classical estimators called, the maximum likelihood, Anderson–Darling, least-squares, Cramér-von Mises, and weighted least squares estimators. The performance of the proposed estimators is explored by detailed simulation results. The flexibility of the OLLLW distribution is studied by two real data sets from medicine and engineering sciences, showing that its capability to fit the data effectively than the Weibull, Fréchet Weibull, transmuted Weibull, gamma Weibull, transmuted exponentiated Weibull, and modified Weibull distributions.

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1. Introduction

Recently, many researchers proposed different ways for generating new extended distributions to enhance and increase its capability to model diverse lifetime data which have a high degree of skewness and kurtosis. These extended distributions can provide more flexibility in modeling several data sets in practice. The exponential, generalized exponential, Weibull, and gamma distributions are adopted for modeling monotonic hazard rates over other models. However, these models are not reasonable for modeling non-monotone hazard rates such as the unimodal and bathtub-shaped hazard rates. The distributions that provide bathtub or/and upside-down bathtub failure rates are very useful in survival and reliability analysis (Kotz et al., 2000).

Several models were proposed to model real lifetime data such as the Weibull distribution as one of the most important distributions for this purpose. Further, the Weibull distribution is adopted

to generate families of distributions such as the Weibull-G family Bourguignon et al. (2014), exponentiated Weibull-H family Cordeiro et al. (2017), extended odd Weibull-G family Alizadeh et al. (2018), Weibull Marshall-Olkin family Korkmaz et al. (2019), and new extended-X family by Zichuan et al. (2020), and new Weibull-X family by Ahmad et al. (2018), among others.

Recently many authors proposed several extended versions and generalizations of the Weibull distribution to increase its flexibility. For example, exponentiated-Weibull by Mudholkar and Srivastava (1993), extended-Weibull by Marshall and Olkin (1997), modified Weibull by Xie et al. (2002), beta-Weibull by Lee et al. (2007), Kumaraswamy-Weibull by Cordeiro et al. (2010), truncated-Weibull by Zhang and Xie (2011), new extended-Weibull by Peng et al. (2014), alpha logarithmic transformed-Weibull by Nassar et al. (2018) Nassar et al. (2018), Kumaraswamy complementary Weibull geometric by Afify et al. (2017), Weibull-Weibull by Abouelmagd et al. (2017), odd log-logistic exponentiated-Weibull by Afify et al. (2018), alpha power exponentiated-Weibull by Mead et al. (2019), odd Lomax-Weibull by Cordeiro et al. (2019), new extended Weibull by Elbatal et al. (2019), Marshall-Olkin power generalized Weibull by Afify et al. (2020), odd Dagum-Weibull by Afify et al. (2020), extended inverse Weibull by Alkarni et al. (2020), logarithmic transformed Weibull by Nassar et al. (2020), and arcsine exponentiated-Weibull distributions by He et al. (2020).

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In this article, we propose a new extension of the Weibull distribution, referred to as the log–logistic Lindley–Weibull (OLLLW) distribution which contains the log–logistic Lindley-exponential and log–logistic Lindley-Rayleigh distributions as special cases. The OLLLW distribution can be considered as a suitable model for modeling skewed data encountered in different applied fields such as biomedical studies, public health, engineering, and survival and reliability analysis. The OLLLW density can be viewed as a mixture of Weibull densities. The OLLLW distribution exhibits monotone, increasing and decreasing, as well as non-monotone, unimodal and bathtub, hazard rates which makes it a superior to other lifetime distributions, which exhibit only increasing or/and decreasing, constant hazard rates. The OLLLW distribution outperforms many common lifetime distributions with respect to real data examples from medicine and engineering sciences.

Furthermore, we study how different classical estimators of the OLLLW parameters perform for several sample sizes and several parameter combinations. The studied classical estimation methods include the maximum likelihood estimators (MLE), least-squares estimators (LSE), Anderson Darling estimators (ADE), weighted least squares estimators (WLSE), and Cramér–von Mises estimators (CVME). The statistical literature contains comprehensive comparisons of different estimation methods for many distributions, such as the quasi xgamma-geometric [Sen et al. \(2019\)](#), Weibull Marshall-Olkin Lindley [Afify et al. \(2020\)](#), odd exponentiated half-logistic exponential [Aldahlan and Afify \(2020\)](#), alpha power exponentiated exponential [Afify et al. \(2020\)](#), generalized Ramos-Louzada [Al-Mofleh et al. \(2020\)](#), and logarithmic transformed Weibull [Nassar et al. \(2020\)](#) distributions.

The OLLLW model is constructed using the odd log–logistic Lindley-G (OLLLi-G) family ([Alizadeh et al. Alizadeh et al. \(2020\)](#)) which is defined by the following cumulative distribution function (CDF)

$$F(x) = \frac{\left\{1 - \left(1 + \frac{\theta}{\theta+1} \frac{G(x; \varphi)}{G(x; \varphi)}\right) e^{-\frac{\theta G(x; \varphi)}{G(x; \varphi)}}\right\}^\alpha}{\left\{1 - \left(1 + \frac{\theta}{\theta+1} \frac{G(x; \varphi)}{G(x; \varphi)}\right) e^{-\frac{\theta G(x; \varphi)}{G(x; \varphi)}}\right\}^\alpha + \left\{\left(1 + \frac{\theta}{\theta+1} \frac{G(x; \varphi)}{G(x; \varphi)}\right) e^{-\frac{\theta G(x; \varphi)}{G(x; \varphi)}}\right\}^{\alpha}}, \quad (1)$$

where $G(x; \varphi)$ denotes a baseline CDF with parameters vector φ , $\alpha > 0$ and $\theta > 0$.

The corresponding probability density function (PDF) of (1) is

$$f(x; \alpha, \theta, \varphi) = \frac{\frac{\alpha \theta^2 g(x; \varphi)}{(1+\theta)\overline{G}(x; \varphi)^3} e^{\frac{\theta G(x; \varphi)}{G(x; \varphi)}} \left(1 + \frac{\theta}{\theta+1} \frac{G(x; \varphi)}{G(x; \varphi)}\right)^{\alpha-1} \left\{1 - \left(1 + \frac{\theta}{\theta+1} \frac{G(x; \varphi)}{G(x; \varphi)}\right) e^{-\frac{\theta G(x; \varphi)}{G(x; \varphi)}}\right\}^{\alpha-1}}{\left(\left\{1 - \left(1 + \frac{\theta}{\theta+1} \frac{G(x; \varphi)}{G(x; \varphi)}\right) e^{-\frac{\theta G(x; \varphi)}{G(x; \varphi)}}\right\}^\alpha + \left\{\left(1 + \frac{\theta}{\theta+1} \frac{G(x; \varphi)}{G(x; \varphi)}\right) e^{-\frac{\theta G(x; \varphi)}{G(x; \varphi)}}\right\}^\alpha\right)^2}, \quad (2)$$

where $g(x; \varphi)$ is a baseline density with a vector of parameters φ .

The paper is organized as follows. In Section 2, we introduce the OLLLW distribution. The properties of the OLLLW distribution are discussed Section 3. Section 4 presents five methods of estimation. In Section 5, detailed simulation results are presented to compare the performance of these estimation methods. In Section 6, the importance of the OLLLW distribution is discussed by two real data sets. Finally, some concluding remarks are given in Section 7.

2. The OLLLW distribution

The CDF of the OLLLW distribution follows, by inserting the CDF and survival function (SF) of the Weibull distribution in Eq. (1), for $x > 0$ and $\lambda, \beta, \alpha, \theta > 0$ as

$$F(x) = \left(\left\{ \frac{e^{\theta-\theta e^{(\lambda x)^\beta}} [\theta e^{(\lambda x)^\beta} + 1]}{\theta + 1} \right\}^\alpha \left\{ 1 - \frac{e^{\theta-\theta e^{(\lambda x)^\beta}} [\theta e^{(\lambda x)^\beta} + 1]}{\theta + 1} \right\}^{-\alpha} + 1 \right)^{-1}. \quad (3)$$

The PDF of the OLLLW model takes the form

$$f(x) = \frac{\alpha \beta \theta^2 (\theta + 1)^{1-2\alpha} \lambda^\beta x^{\beta-1} \left(\left[\theta e^{(\lambda x)^\beta} + 1 \right] \left\{ (\theta + 1) e^{\theta e^{(\lambda x)^\beta}} - e^\theta [\theta e^{(\lambda x)^\beta} + 1] \right\} \right)^{\alpha-1}}{\left(\left\{ \frac{e^{\theta-\theta e^{(\lambda x)^\beta}} [\theta e^{(\lambda x)^\beta} + 1]}{\theta + 1} \right\}^\alpha + \left\{ 1 - \frac{e^{\theta-\theta e^{(\lambda x)^\beta}} [\theta e^{(\lambda x)^\beta} + 1]}{\theta + 1} \right\}^\alpha \right)^2 e^{-[\alpha \theta + (\theta - 2\alpha \theta) e^{(\lambda x)^\beta} + 2(\lambda x)^\beta]}}. \quad (4)$$

The SF and hazard rate function (HRF) of the OLLLW distribution have the forms

$$S(x) = \frac{1}{e^{-\alpha \theta \left(\frac{(\theta+1)e^{\theta e^{(\lambda x)^\beta}} + 1}{\theta e^{(\lambda x)^\beta} + 1} - e^\theta \right)^\alpha} + 1}, \quad (5)$$

$$h(x) = \alpha \beta \theta^2 (\theta + 1)^{1-2\alpha} \lambda^\beta x^{\beta-1} \left(\left[\theta e^{(\lambda x)^\beta} + 1 \right] \left\{ (\theta + 1) e^{\theta e^{(\lambda x)^\beta}} - e^\theta [\theta e^{(\lambda x)^\beta} + 1] \right\} \right)^{\alpha-1} \quad (6)$$

$$\times \frac{\left\{ e^{-\alpha \theta \left(\frac{(\theta+1)e^{\theta e^{(\lambda x)^\beta}} + 1}{\theta e^{(\lambda x)^\beta} + 1} - e^\theta \right)^\alpha} + 1 \right\} e^{[\alpha \theta + (\theta - 2\alpha \theta) e^{(\lambda x)^\beta} + 2(\lambda x)^\beta]}}{\left(\left\{ \frac{e^{\theta-\theta e^{(\lambda x)^\beta}} [\theta e^{(\lambda x)^\beta} + 1]}{\theta + 1} \right\}^\alpha + \left\{ 1 - \frac{e^{\theta-\theta e^{(\lambda x)^\beta}} [\theta e^{(\lambda x)^\beta} + 1]}{\theta + 1} \right\}^\alpha \right)^2}. \quad (7)$$

Plots of the PDF and HRF of the OLLLW distribution are shown in [Figs. 1 and 2](#), respectively. These plots show that the OLLLW distribution can provide left-skewed, symmetrical, right-skewed or reversed-J shaped densities, and its HRF exhibits bathtub, decreasing, unimodel, increasing, J shaped or reversed-J shaped hazard rates.

3. Mathematical Properties

3.1. Linear Representation

By using Eq. (10) introduced by Alizadeh et al. [Alizadeh et al. \(2020\)](#), we have

$$f(x) = \sum_{k,s,m=0}^{\infty} \sum_{l=0}^r \sum_{r=0}^k w_{k,r,s,l,m} (s+l+m+1) g(x) G(x)^{s+l+m},$$

then by using the binomial expansion $(1-z)^b = \sum_{j=0}^{\infty} (-1)^j \binom{b}{j} z^j$, we have

$$\begin{aligned} f(x) &= \sum_{k,s,m=0}^{\infty} \sum_{l=0}^r \sum_{r=0}^k w_{k,r,s,l,m} (s+l+m+1) \beta \lambda^\beta x^{\beta-1} e^{-(\lambda x)^\beta} \left[1 - e^{-(\lambda x)^\beta} \right]^{s+l+m} \\ &= \sum_{k,s,m=0}^{\infty} \sum_{l=0}^r \sum_{r=0}^k \sum_{j=0}^{\infty} (-1)^j \binom{s+l+m}{j} w_{k,r,s,l,m} (s+l+m+1) \beta \lambda^\beta x^{\beta-1} e^{-\lambda^\beta (j+1)x^\beta} \\ &= \sum_{j=0}^{\infty} \Phi_j q_j(x), \end{aligned}$$

where $\Phi_j = \Phi(k, s, m, l, r) = \sum_{k,s,m=0}^{\infty} \sum_{l=0}^r \sum_{r=0}^k (-1)^j \binom{s+l+m+1}{j+1} w_{k,r,s,l,m}$ and $q_j(x)$ is the PDF of Weibull distribution with shape parameter β and scale parameter $\lambda(j+1)^{\frac{1}{\beta}}$.

3.2. Moments

The r th moments of the OLLLW distribution has the form

$$\mu'_r = E(X^r) = \int_0^{\infty} x^r f(x) dx = \sum_{j=0}^{\infty} \Phi_{k,s,m,l,j} \left[\lambda (j+1)^{\frac{1}{\beta}} \right]^{-r} \Gamma \left(1 + \frac{r}{\beta} \right).$$

Setting $r = 1, 2, 3$, and 4 , respectively, we obtain the first four moments about the origin of the OLLLW distribution.

The n th central moment of X , say μ_n , follows as

$$\mu_n = E(x - \mu)^n = \sum_{k=0}^{\infty} (-1)^k \binom{n}{k} \mu_1^k \mu'_{n-k}.$$

The cumulants (k_n) of X can be obtained as the following

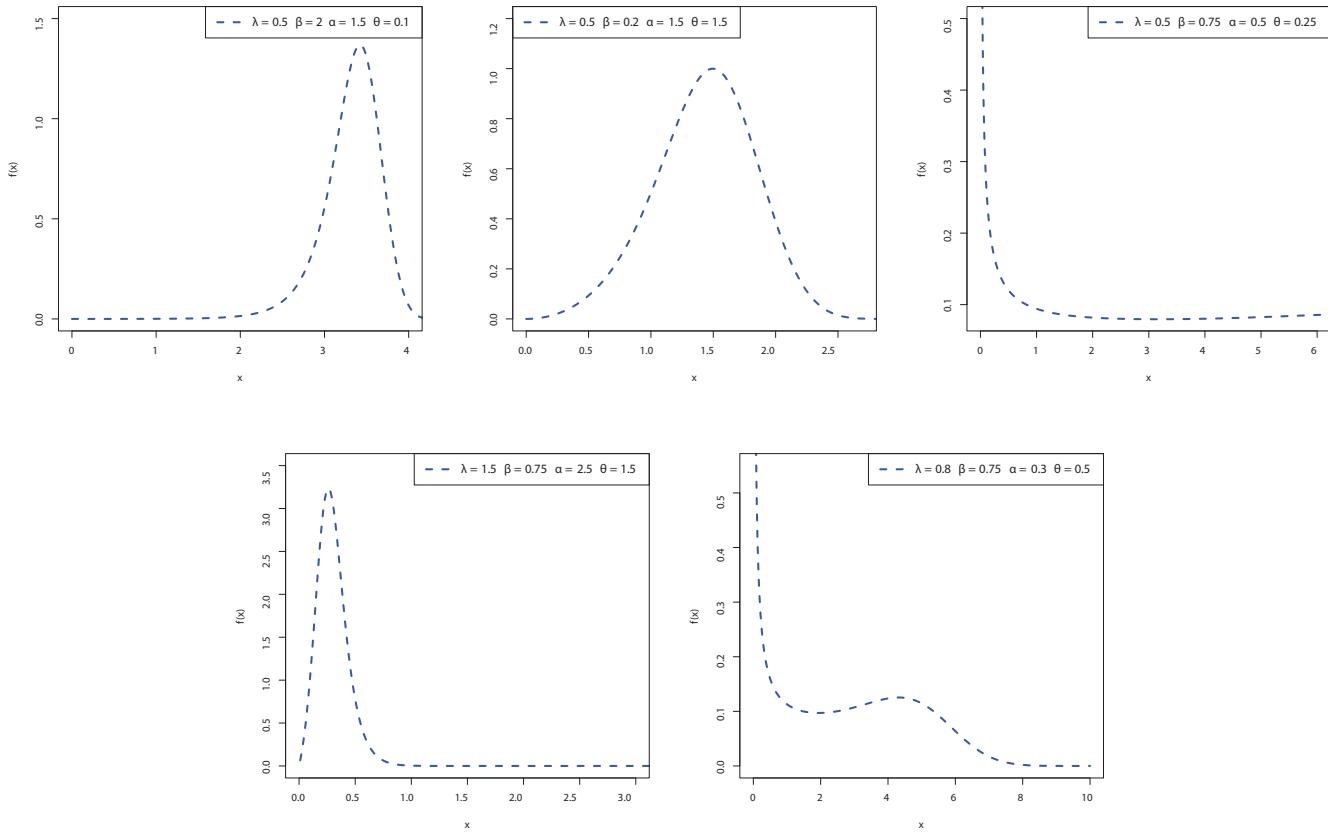


Fig. 1. Some possible shapes of the PDF of the OLLLW distribution.

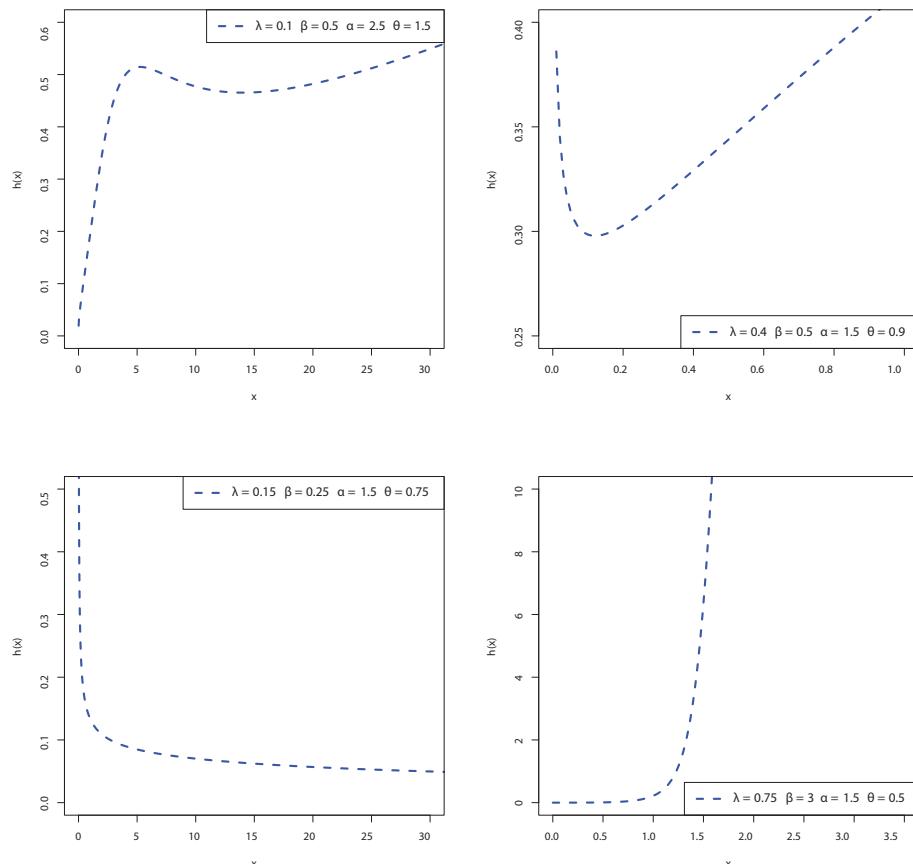


Fig. 2. Some possible shapes of the HRF of the OLLLW distribution.

$$k_n = \mu'_n - \sum_{r=0}^{n-1} \binom{n-1}{r-1} k_r \mu'_{n-r}.$$

3.3. Incomplete Moments

The s th incomplete moment of Weibull distribution is

$$I_s = \gamma^{-s} B\left(1 + \frac{s}{\beta}, (\gamma t)^\beta\right),$$

where $B(s, t)$ is lower incomplete gamma function.

The s th incomplete moments of the OLLLW distribution is

$$\begin{aligned} \Psi_s(t) &= \int_0^t x^s f(x) dx \\ &= \sum_{j=0}^{\infty} \Phi_{k,s,m,l,r,j} [\lambda^\beta (j+1)]^{-s} B\left(1 + \frac{s}{\beta}, (\lambda(j+1)^{\frac{1}{\beta}} t)^\beta\right). \end{aligned}$$

The mean residual life of the OLLLW distribution is

$$m_x(t) = \frac{1 - \Psi_1(t)}{R(t) - t}.$$

The mean inactivity time of the OLLLW distribution is

$$M_x(t) = t - \frac{\Psi_1(t)}{F(t)}.$$

3.4. Generating Function

the moment generating function (MGF) of Weibull distribution is given by

$$M(t) = \int_0^\infty e^{tx} f(x) dx = \beta \gamma^\beta \int_0^\infty e^{tx} x^{\beta-1} e^{-(\gamma x)^\beta} dx,$$

By expanding the first exponential and calculating the integral, we can write

$$M(t) = \sum_{m=0}^{\infty} \frac{(t/\gamma)^m}{m!} \Gamma(\beta + m/\beta).$$

Consider the Wright generalized hyper geometric function which is defined by

$$\Omega_q \left[\begin{matrix} (\gamma_1, A_1), \dots, (\gamma_p, A_p) \\ (\beta_1, B_1), \dots, (\beta_q, B_q) \end{matrix}; x \right] = \sum_{n=0}^{\infty} \frac{\prod_{j=1}^p \Gamma(\gamma_j + A_j n)}{\prod_{j=1}^q \Gamma(\beta_j + B_j n) n!}.$$

Hence, we can write the MGF as

$$M(t) = {}_1\Omega_0 \left[\begin{matrix} (1, -\beta^{-1}) \\ - \end{matrix}; \frac{t}{\gamma} \right].$$

Hence, the MGF of the OLLLW model is given by

$$M(t) = \sum_{j=0}^{\infty} \Phi_{k,s,m,l,r,j} {}_1\Omega_0 \left[\begin{matrix} (1, -\beta^{-1}) \\ - \end{matrix}; \frac{t}{\lambda(j+1)^{\frac{1}{\beta}}} \right].$$

Characteristic function of the OLLLW distribution can be determined from the last equation by setting $t = it$.

3.5. Order Statistics

The PDF and CDF of the i th order statistic for the OLLLW distribution are

$$\begin{aligned} f_{i:n}(x) &= \frac{f(x)}{B(i,n-i+1)} \sum_{h=0}^{\infty} (-1)^h \binom{n-i}{h} F^{h+i-1}(x) \\ &= \frac{\sum_{j=0}^{\infty} \Phi_{k,s,m,l,r,j} q_j(x)}{B(i,n-i+1)} \sum_{h=0}^{\infty} (-1)^h \binom{n-i}{h} \\ &\times \left(\left\{ \frac{e^{\theta-\theta e^{(\lambda x)^\beta}} [\theta e^{(\lambda x)^\beta} + 1]}{\theta + 1} \right\}^\alpha \left\{ 1 - \frac{e^{\theta-\theta e^{(\lambda x)^\beta}} [\theta e^{(\lambda x)^\beta} + 1]}{\theta + 1} \right\}^{-\alpha} + 1 \right)^{-(h+i-1)}. \end{aligned}$$

where $B(.,.)$ is beta function.

$$\begin{aligned} F_{i:n}(x) &= \sum_{r=i}^n \binom{n}{r} (F(x))^r (1 - F(x))^{n-r} \\ &= \frac{\Gamma(n+1)}{\Gamma(-i+n+1)} \left\{ e^{x\theta} \left[\frac{(\theta+1)e^{\theta e^{(\lambda x)^\beta}}}{\theta e^{(\lambda x)^\beta} + 1} - e^\theta \right]^{-\alpha} + 1 \right\}^{-i} \\ &\times \left\{ e^{x\theta} \left[\frac{(\theta+1)e^{\theta e^{(\lambda x)^\beta}}}{\theta e^{(\lambda x)^\beta} + 1} - e^\theta \right]^\alpha + 1 \right\}^{-n+i} {}_2\tilde{\Omega}_1 \\ &\times \left(1, i-n; i+1; - \left\{ \frac{e^{[-1+e^{(\lambda x)^\beta}]_\theta} (\theta+1)}{e^{(\lambda x)^\beta} \theta + 1} - 1 \right\}^\alpha \right), \end{aligned}$$

where ${}_2\tilde{\Omega}_1 \left(1, i-n; i+1; - \left\{ \frac{e^{[-1+e^{(\lambda x)^\beta}]_\theta} (\theta+1)}{e^{(\lambda x)^\beta} \theta + 1} - 1 \right\}^\alpha \right)$ is a hyper geometric regularized function.

4. Methods of Estimation

In this section, we discuss the estimation of the OLLLW parameters by different estimators including the MLE, ADE, CVME, LSE and WLSE.

4.1. Maximum Likelihood Estimation

Let x_1, x_2, \dots, x_n be a random sample from the PDF (4), then the log-likelihood function reduces to

$$\begin{aligned} L &= (\beta - 1) \sum_{i=1}^n \log(x_i) + n \log \left(\alpha \beta \theta^2 (\theta + 1)^{1-2\alpha} \lambda^\beta \right) \\ &+ \sum_{i=1}^n \left(\alpha \theta + (\theta - 2\alpha \theta) e^{(\lambda x_i)^\beta} + 2(\lambda x_i)^\beta \right) + (\alpha - 1) \sum_{i=1}^n \log \left(\theta e^{(\lambda x_i)^\beta} + 1 \right) \\ &+ (\alpha - 1) \sum_{i=1}^n \log \left((\theta + 1) e^{\theta e^{(\lambda x_i)^\beta}} - e^\theta (\theta e^{(\lambda x_i)^\beta} + 1) \right) \\ &- 2 \sum_{i=1}^n \log \left(\left(\frac{e^{\theta-\theta e^{(\lambda x_i)^\beta}} (\theta e^{(\lambda x_i)^\beta} + 1)}{\theta + 1} \right)^\alpha + \left(1 - \frac{e^{\theta-\theta e^{(\lambda x_i)^\beta}} (\theta e^{(\lambda x_i)^\beta} + 1)}{\theta + 1} \right)^\alpha \right). \end{aligned} \quad (8)$$

By differentiating Eq. (8) with respect to α, β and λ , respectively, equating to zero, and solving the previous equations, we obtain estimators of the OLLLW parameters by the MLE.

4.2. Ordinary Least-Squares and Weighted Least-Squares Estimators

Let $x_{1:n}, x_{2:n}, \dots, x_{2:n}$ be the order statistics of a random sample of size n from the OLLLW distribution. Hence, we have the OLSE of the OLLLW parameters by minimizing the following equation:

$$O = \sum_{i=1}^n \left[F(x_{i:n}) - \frac{i}{n+1} \right]^2,$$

The OLSE of the OLLLW parameters can also be obtained by solving the following nonlinear equations:

Table 1Simulation results of the AEs, ABBs, MSEs and MREs of the OLLLW distribution for ($\lambda = 0.5, \beta = 0.5, \alpha = 0.5, \theta = 0.5$).

n	Est.	Est. Par.	MLE	ADE	CVME	LSE	WLSE
20	AEs	$\hat{\lambda}$	0.37223	0.51074	0.55029	0.63779	0.97212
		$\hat{\beta}$	0.65793	0.54499	0.57692	0.53378	0.57701
		$\hat{\alpha}$	0.52270	0.56383	0.54524	0.52665	0.53563
		$\hat{\theta}$	0.71352	0.62252	0.62519	0.57207	0.83341
	ABBs	$\hat{\lambda}$	0.21138	0.23141	0.35266	0.33645	0.76128
		$\hat{\beta}$	0.18493	0.11925	0.17748	0.16140	0.21868
		$\hat{\alpha}$	0.11798	0.13697	0.19243	0.18946	0.20673
		$\hat{\theta}$	0.27923	0.24677	0.33873	0.31430	0.63312
	MSEs	$\hat{\lambda}$	0.06868	0.08729	0.13617	0.13933	0.91268
		$\hat{\beta}$	0.06639	0.03076	0.05331	0.04326	0.10101
		$\hat{\alpha}$	0.03056	0.04216	0.06004	0.05726	0.10237
		$\hat{\theta}$	0.12409	0.10255	0.14956	0.12441	0.67105
	MREs	$\hat{\lambda}$	0.42276	0.46282	0.64531	0.67290	1.52256
		$\hat{\beta}$	0.36986	0.23850	0.35496	0.32281	0.43736
		$\hat{\alpha}$	0.23596	0.27395	0.38486	0.37893	0.41345
		$\hat{\theta}$	0.55847	0.49354	0.65374	0.62175	1.30622
30	AEs	$\hat{\lambda}$	0.38296	0.49641	0.53979	0.60426	0.88481
		$\hat{\beta}$	0.61424	0.54051	0.56506	0.53475	0.57499
		$\hat{\alpha}$	0.52219	0.54683	0.53903	0.52592	0.50547
		$\hat{\theta}$	0.70172	0.62755	0.65368	0.60927	0.86915
	ABBs	$\hat{\lambda}$	0.18335	0.20960	0.31309	0.33148	0.70540
		$\hat{\beta}$	0.13172	0.09440	0.15693	0.14589	0.18752
		$\hat{\alpha}$	0.08564	0.09803	0.16262	0.15973	0.14469
		$\hat{\theta}$	0.24935	0.22346	0.32364	0.29692	0.60557
	MSEs	$\hat{\lambda}$	0.05706	0.07850	0.12006	0.13496	0.81529
		$\hat{\beta}$	0.04125	0.02244	0.04219	0.03564	0.06127
		$\hat{\alpha}$	0.01898	0.02582	0.04422	0.04220	0.04033
		$\hat{\theta}$	0.11310	0.09313	0.13341	0.11700	0.66178
	MREs	$\hat{\lambda}$	0.36669	0.41920	0.62618	0.66296	1.41080
		$\hat{\beta}$	0.26343	0.18879	0.31387	0.29178	0.37503
		$\hat{\alpha}$	0.17128	0.19607	0.32524	0.31945	0.28937
		$\hat{\theta}$	0.49869	0.44691	0.64727	0.59383	1.21114
50	AEs	$\hat{\lambda}$	0.39306	0.47858	0.53531	0.58913	0.77840
		$\hat{\beta}$	0.58001	0.53397	0.55464	0.53025	0.58495
		$\hat{\alpha}$	0.51806	0.52813	0.52119	0.51285	0.49067
		$\hat{\theta}$	0.67093	0.60877	0.65169	0.61320	0.88552
	ABBs	$\hat{\lambda}$	0.13175	0.14864	0.30883	0.32435	0.61053
		$\hat{\beta}$	0.09040	0.06684	0.12809	0.11680	0.17510
		$\hat{\alpha}$	0.05425	0.06016	0.11940	0.11728	0.10472
		$\hat{\theta}$	0.19068	0.16782	0.29954	0.28470	0.58576
	MSEs	$\hat{\lambda}$	0.03873	0.05339	0.11753	0.12927	0.65511
		$\hat{\beta}$	0.02578	0.01450	0.02843	0.02341	0.05181
		$\hat{\alpha}$	0.00972	0.01217	0.02503	0.02380	0.01939
		$\hat{\theta}$	0.08718	0.07044	0.12238	0.11163	0.64481
	MREs	$\hat{\lambda}$	0.26351	0.29727	0.61767	0.64871	1.22106
		$\hat{\beta}$	0.18079	0.13367	0.25617	0.23359	0.35020
		$\hat{\alpha}$	0.10850	0.12032	0.23880	0.23455	0.20943
		$\hat{\theta}$	0.38135	0.33564	0.59908	0.56940	1.17151
100	AEs	$\hat{\lambda}$	0.43693	0.47596	0.53570	0.57791	0.65754
		$\hat{\beta}$	0.53636	0.51950	0.55015	0.53402	0.56879
		$\hat{\alpha}$	0.50769	0.50803	0.49923	0.49549	0.48535
		$\hat{\theta}$	0.59155	0.55544	0.62674	0.59622	0.76327
	ABBs	$\hat{\lambda}$	0.06698	0.06741	0.29317	0.30538	0.45102
		$\hat{\beta}$	0.03795	0.02797	0.10406	0.09751	0.13486
		$\hat{\alpha}$	0.01845	0.01885	0.07748	0.07719	0.07010
		$\hat{\theta}$	0.09765	0.08068	0.26707	0.25408	0.42273
	MSEs	$\hat{\lambda}$	0.01846	0.02070	0.10874	0.11822	0.40129
		$\hat{\beta}$	0.00777	0.00488	0.01892	0.01666	0.03201
		$\hat{\alpha}$	0.00206	0.00239	0.01055	0.01024	0.00777
		$\hat{\theta}$	0.04373	0.03293	0.10275	0.09403	0.37701
	MREs	$\hat{\lambda}$	0.13395	0.13483	0.58634	0.61075	0.90205
		$\hat{\beta}$	0.07591	0.05593	0.20813	0.19502	0.26972
		$\hat{\alpha}$	0.03689	0.03770	0.15496	0.15439	0.14019
		$\hat{\theta}$	0.19529	0.16136	0.53414	0.50816	0.84545
200	AEs	$\hat{\lambda}$	0.48270	0.48961	0.53164	0.55683	0.57714
		$\hat{\beta}$	0.50990	0.50641	0.54171	0.53223	0.55620

(continued on next page)

Table 1 (continued)

n	Est.	Est. Par.	MLE	ADE	CVME	LSE	WLSE
400	ABBs	$\hat{\alpha}$	0.50063	0.50052	0.49494	0.49322	0.48836
		$\hat{\theta}$	0.52459	0.51536	0.60689	0.58806	0.69250
		$\hat{\lambda}$	0.01730	0.01583	0.27122	0.27779	0.35175
		$\hat{\beta}$	0.00990	0.00704	0.09254	0.08849	0.10685
	MSEs	$\hat{\alpha}$	0.00308	0.00297	0.05785	0.05760	0.05027
		$\hat{\theta}$	0.02478	0.01823	0.22915	0.22056	0.31578
		$\hat{\lambda}$	0.00487	0.00429	0.09690	0.10238	0.26557
		$\hat{\beta}$	0.00188	0.00109	0.01325	0.01201	0.01897
	MREs	$\hat{\alpha}$	0.00023	0.00025	0.00538	0.00530	0.00398
		$\hat{\theta}$	0.01056	0.00703	0.07859	0.07338	0.21204
		$\hat{\lambda}$	0.03459	0.03166	0.54245	0.55558	0.70351
		$\hat{\beta}$	0.01980	0.01409	0.18507	0.17697	0.21371
	AEs	$\hat{\alpha}$	0.00617	0.00593	0.11570	0.11520	0.10053
		$\hat{\theta}$	0.04957	0.03646	0.45830	0.44111	0.63155
		$\hat{\lambda}$	0.49937	0.49939	0.53569	0.55121	0.49936
		$\hat{\beta}$	0.50044	0.50036	0.52805	0.52258	0.54322
	ABBs	$\hat{\alpha}$	0.49993	0.49995	0.49392	0.49310	0.49304
		$\hat{\theta}$	0.50100	0.50095	0.57092	0.55951	0.62871
		$\hat{\lambda}$	0.00063	0.00061	0.23134	0.23418	0.23527
		$\hat{\beta}$	0.00044	0.00036	0.07262	0.07058	0.07707
	MSEs	$\hat{\alpha}$	0.00007	0.00005	0.04078	0.04081	0.03632
		$\hat{\theta}$	0.00100	0.00095	0.18371	0.17947	0.21208
		$\hat{\lambda}$	0.00020	0.00018	0.07736	0.07994	0.11972
		$\hat{\beta}$	0.00010	0.00007	0.00806	0.00759	0.01010
	MREs	$\hat{\alpha}$	0.00000	0.00000	0.00271	0.00271	0.00209
		$\hat{\theta}$	0.00050	0.00046	0.05470	0.05205	0.09328
		$\hat{\lambda}$	0.00125	0.00121	0.46268	0.46836	0.47055
		$\hat{\beta}$	0.00088	0.00072	0.14525	0.14117	0.15415
		$\hat{\alpha}$	0.00014	0.00010	0.08156	0.08161	0.07264
		$\hat{\theta}$	0.00200	0.00191	0.36742	0.35894	0.42417

$$\sum_{i=1}^n \left\{ F(x_{i:n}) - \frac{i}{n+1} \right\} \Delta_s(x_{i:n}) = 0, \quad s = 1, 2, 3, 4,$$

where $\Delta_1(x_{i:n}) = \frac{\partial}{\partial \alpha} F(x_{i:n})$, $\Delta_2(x_{i:n}) = \frac{\partial}{\partial \beta} F(x_{i:n})$, $\Delta_3(x_{i:n}) = \frac{\partial}{\partial \theta} F(x_{i:n})$, and $\Delta_4(x_{i:n}) = \frac{\partial}{\partial \lambda} F(x_{i:n})$.

The WLSE of the OLLLW parameters can be calculated by minimizing the following equation:

$$W = \sum_{i=1}^n \frac{(n+1)^2(n+2)}{i(n-i+1)} \left[F(x_{i:n}) - \frac{i}{n+1} \right]^2.$$

Furthermore, the WLSE of the OLLLW parameters follow by solving the following nonlinear equations:

$$\sum_{i=1}^n \frac{(n+1)^2(n+2)}{i(n-i+1)} \left[F(x_{i:n}) - \frac{i}{n+1} \right] \Delta_s(x_{i:n}) = 0.$$

4.3. Anderson–Darling Estimation

The ADE of the OLLLW parameters are obtained by minimizing the following equation:

$$A = -n - \frac{1}{n} \sum_{i=1}^n (2i-1) [\log F(x_{i:n}) + \log S(x_{i:n})].$$

The ADE can also be calculated by solving the following nonlinear equations:

$$\sum_{i=1}^n (2i-1) \left[\frac{\Delta_s(x_{i:n})}{F(x_{i:n})} - \frac{\Delta_s(x_{n+1-i:n})}{S(x_{n+1-i:n})} \right] = 0.$$

4.4. Cramér–von Mises Estimation

The CVME of OLLLW parameters are obtained by minimizing the following equation:

$$CV = \frac{1}{12n} + \sum_{i=1}^n \left[F(x_{i:n}) - \frac{2i-1}{2n} \right]^2,$$

or by solving the following nonlinear equations

$$\sum_{i=1}^n \left\{ F(x_{i:n}) - \frac{2i-1}{2n} \right\} \Delta_s(x_{i:n}) = 0.$$

5. Simulation Results

This section is devoted to exploring the behavior of the previous estimation methods for estimating the OLLLW parameters based on a detailed simulation study. Some different sample sizes, $n = \{20, 30, 50, 100, 200, 400\}$, are considered along with some different parametric values of λ, β, α and θ , where $\lambda = \{0.5, 1.5, 3.5\}$, $\beta = \{0.25, 0.5, 0.75, 1.75\}$, $\alpha = \{0.5, 1.5, 2.5\}$ and $\theta \in \{0.5, 1.5, 2.5\}$. The simulations were repeated $N = 2000$ times, where the random samples are generated from the OLLLW distribution to determine the following measures such as, average estimates (AEs) with their associated mean square error (MSEs), average absolute biases (ABBs), and mean relative errors of the estimates (MREs) for all studied sample sizes and parameters combinations using the statistical R software®.

The ABBs, MSEs, and MREs are given by:

Table 2Simulation results of the AEs, ABBs, MSEs and MREs of the OLLLW distribution for ($\lambda = 0.5$, $\beta = 0.25$, $\alpha = 2.5$, $\theta = 1.5$).

n	Est.	Est. Par.	MLE	ADE	CVME	LSE	WLSE
20	AEs	$\hat{\lambda}$	0.60662	0.58062	0.60480	0.59165	0.60372
		$\hat{\beta}$	0.38281	0.33324	0.36076	0.33855	0.32811
		$\hat{\alpha}$	2.12735	2.30821	2.25169	2.23353	2.26736
		$\hat{\theta}$	1.94566	1.84618	1.88502	1.83063	1.81747
	ABBs	$\hat{\lambda}$	0.35295	0.37936	0.37244	0.38723	0.40001
		$\hat{\beta}$	0.14585	0.10963	0.13446	0.12526	0.11310
		$\hat{\alpha}$	0.72387	0.66685	0.71679	0.74239	0.69585
		$\hat{\theta}$	0.46381	0.41852	0.42103	0.39541	0.42251
	MSEs	$\hat{\lambda}$	0.14965	0.16473	0.16006	0.17014	0.17800
		$\hat{\beta}$	0.03855	0.02427	0.03388	0.02976	0.02490
		$\hat{\alpha}$	0.68675	0.58697	0.66203	0.71097	0.63339
		$\hat{\theta}$	0.22836	0.19951	0.20122	0.18486	0.19945
	MREs	$\hat{\lambda}$	0.70590	0.75871	0.74488	0.77446	0.80003
		$\hat{\beta}$	0.58338	0.43851	0.53783	0.50104	0.45240
		$\hat{\alpha}$	0.28955	0.26674	0.28672	0.29696	0.27834
		$\hat{\theta}$	0.30920	0.27901	0.28069	0.26361	0.28168
30	AEs	$\hat{\lambda}$	0.57553	0.56142	0.56486	0.56092	0.58151
		$\hat{\beta}$	0.35775	0.31403	0.34303	0.32729	0.31773
		$\hat{\alpha}$	2.18043	2.36659	2.29935	2.28934	2.31682
		$\hat{\theta}$	1.92034	1.81370	1.88002	1.83660	1.81784
	ABBs	$\hat{\lambda}$	0.33721	0.37049	0.36164	0.37358	0.39361
		$\hat{\beta}$	0.12234	0.09084	0.11753	0.11128	0.09965
		$\hat{\alpha}$	0.68138	0.60885	0.68481	0.69667	0.65656
		$\hat{\theta}$	0.45378	0.40992	0.41350	0.39233	0.42213
	MSEs	$\hat{\lambda}$	0.13980	0.15835	0.15327	0.16062	0.17102
		$\hat{\beta}$	0.02688	0.01679	0.02647	0.02376	0.01940
		$\hat{\alpha}$	0.59416	0.48601	0.59620	0.62019	0.54451
		$\hat{\theta}$	0.22160	0.19564	0.19844	0.18579	0.19915
	MREs	$\hat{\lambda}$	0.67442	0.74098	0.72328	0.74717	0.78722
		$\hat{\beta}$	0.48937	0.36335	0.47010	0.44513	0.39859
		$\hat{\alpha}$	0.27255	0.24354	0.27393	0.27867	0.26262
		$\hat{\theta}$	0.30252	0.27430	0.27567	0.26155	0.28142
50	AEs	$\hat{\lambda}$	0.53156	0.52859	0.54069	0.54574	0.54924
		$\hat{\beta}$	0.33227	0.30273	0.32414	0.31294	0.30297
		$\hat{\alpha}$	2.25327	2.40549	2.33426	2.33746	2.39520
		$\hat{\theta}$	1.88757	1.81115	1.85212	1.81235	1.80534
	ABBs	$\hat{\lambda}$	0.31748	0.35421	0.35998	0.36909	0.38035
		$\hat{\beta}$	0.09605	0.07845	0.09746	0.09285	0.08534
		$\hat{\alpha}$	0.60962	0.57085	0.63606	0.63893	0.60944
		$\hat{\theta}$	0.43523	0.40767	0.39665	0.38037	0.41695
	MSEs	$\hat{\lambda}$	0.12606	0.14828	0.15139	0.15980	0.16289
		$\hat{\beta}$	0.01798	0.01330	0.01927	0.01778	0.01473
		$\hat{\alpha}$	0.48968	0.43010	0.51754	0.52650	0.47179
		$\hat{\theta}$	0.21048	0.19181	0.18834	0.17739	0.19864
	MREs	$\hat{\lambda}$	0.63497	0.70841	0.71968	0.73925	0.76070
		$\hat{\beta}$	0.38418	0.31381	0.38983	0.37138	0.34135
		$\hat{\alpha}$	0.24385	0.22834	0.25442	0.25557	0.24378
		$\hat{\theta}$	0.29015	0.27178	0.26443	0.25358	0.27695
100	AEs	$\hat{\lambda}$	0.49688	0.54568	0.55172	0.56469	0.55990
		$\hat{\beta}$	0.30508	0.28811	0.31615	0.30864	0.29570
		$\hat{\alpha}$	2.33352	2.41247	2.31595	2.32343	2.37830
		$\hat{\theta}$	1.83147	1.74308	1.81700	1.78208	1.76966
	ABBs	$\hat{\lambda}$	0.28503	0.34351	0.35830	0.36792	0.37706
		$\hat{\beta}$	0.06935	0.05944	0.08804	0.08482	0.07109
		$\hat{\alpha}$	0.49696	0.48371	0.60743	0.60602	0.54822
		$\hat{\theta}$	0.40362	0.35723	0.38196	0.36595	0.39950
	MSEs	$\hat{\lambda}$	0.10827	0.14425	0.14930	0.15604	0.16140
		$\hat{\beta}$	0.00991	0.00786	0.01568	0.01461	0.01052
		$\hat{\alpha}$	0.34879	0.32328	0.47267	0.47191	0.38506
		$\hat{\theta}$	0.19202	0.16256	0.17874	0.16881	0.18310
	MREs	$\hat{\lambda}$	0.57006	0.68702	0.71659	0.73584	0.75412
		$\hat{\beta}$	0.27741	0.23775	0.35215	0.33928	0.28435
		$\hat{\alpha}$	0.19878	0.19348	0.24297	0.24241	0.21929
		$\hat{\theta}$	0.26908	0.23816	0.25464	0.24397	0.26633
200	AEs	$\hat{\lambda}$	0.46952	0.53347	0.54130	0.55761	0.55846
		$\hat{\beta}$	0.28388	0.27314	0.30165	0.29681	0.28224
		$\hat{\alpha}$	2.40986	2.45782	2.34591	2.35310	2.41006

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Table 2 (continued)

n	Est.	Est. Par.	MLE	ADE	CVME	LSE	WLSE
400	ABBs	$\hat{\theta}$	1.77638	1.69180	1.78321	1.75370	1.72396
		$\hat{\lambda}$	0.24461	0.29458	0.33986	0.34714	0.35044
		$\hat{\beta}$	0.04812	0.04446	0.07290	0.07104	0.05665
		$\hat{\alpha}$	0.37204	0.38208	0.54866	0.54642	0.46543
	MSEs	$\hat{\theta}$	0.36257	0.32611	0.36465	0.35395	0.38061
		$\hat{\lambda}$	0.08926	0.12005	0.13808	0.14292	0.14542
		$\hat{\beta}$	0.00493	0.00418	0.01053	0.00998	0.00610
		$\hat{\alpha}$	0.21224	0.21573	0.38326	0.38029	0.27856
	MREs	$\hat{\theta}$	0.17022	0.14664	0.16702	0.16069	0.17139
		$\hat{\lambda}$	0.48922	0.58916	0.67972	0.69428	0.70087
		$\hat{\beta}$	0.19249	0.17784	0.29159	0.28417	0.22659
		$\hat{\alpha}$	0.14881	0.15283	0.21946	0.21857	0.18617
	AEs	$\hat{\theta}$	0.24171	0.21740	0.24310	0.23597	0.25374
		$\hat{\lambda}$	0.44691	0.53808	0.55211	0.56510	0.56397
		$\hat{\beta}$	0.27323	0.26307	0.28519	0.28221	0.27214
		$\hat{\alpha}$	2.44607	2.48730	2.39294	2.39861	2.43828
	ABBs	$\hat{\theta}$	1.74363	1.63971	1.71920	1.69882	1.68089
		$\hat{\lambda}$	0.20993	0.26812	0.32984	0.33697	0.33900
		$\hat{\beta}$	0.03557	0.03223	0.05731	0.05616	0.04545
		$\hat{\alpha}$	0.26300	0.28064	0.45825	0.45652	0.37903
	MSEs	$\hat{\theta}$	0.32595	0.28731	0.34307	0.33580	0.36802
		$\hat{\lambda}$	0.07275	0.10981	0.13339	0.13833	0.13961
		$\hat{\beta}$	0.00276	0.00233	0.00633	0.00602	0.00369
		$\hat{\alpha}$	0.12468	0.13408	0.27307	0.27087	0.19427
	MREs	$\hat{\theta}$	0.15112	0.12645	0.15352	0.14879	0.15984
		$\hat{\lambda}$	0.41986	0.53623	0.65968	0.67394	0.67801
		$\hat{\beta}$	0.14226	0.12892	0.22925	0.22464	0.18180
		$\hat{\alpha}$	0.10520	0.11226	0.18330	0.18261	0.15161
		$\hat{\theta}$	0.21730	0.19154	0.22872	0.22387	0.24535

Table 3Simulation results of the AEs, ABBs, MSEs and MREs of the OLLLW distribution for ($\lambda = 3.5$, $\beta = 0.75$, $\alpha = 1.5$, $\theta = 2.5$).

n	Est.	Est. Par.	MLE	ADE	CVME	LSE	WLSE
20	AEs	$\hat{\lambda}$	3.57711	3.49269	3.51955	3.49103	3.50103
		$\hat{\beta}$	0.87859	0.80235	0.82525	0.78589	0.78077
		$\hat{\alpha}$	1.46297	1.53845	1.55203	1.49161	1.52253
		$\hat{\theta}$	3.03006	2.74771	2.81906	2.66580	2.65335
	ABBs	$\hat{\lambda}$	0.49528	0.47334	0.46764	0.46305	0.47471
		$\hat{\beta}$	0.19829	0.18195	0.19001	0.18907	0.18176
		$\hat{\alpha}$	0.34083	0.35374	0.35820	0.36926	0.35542
		$\hat{\theta}$	0.71318	0.59115	0.60660	0.54899	0.57486
	MSEs	$\hat{\lambda}$	0.25322	0.23250	0.22843	0.22574	0.23349
		$\hat{\beta}$	0.04570	0.03996	0.04255	0.04237	0.04020
		$\hat{\alpha}$	0.14741	0.15747	0.16112	0.17123	0.16068
		$\hat{\theta}$	0.82908	0.57485	0.63813	0.50113	0.52679
	MREs	$\hat{\lambda}$	0.14579	0.13524	0.13361	0.13230	0.13563
		$\hat{\beta}$	0.26439	0.24261	0.25335	0.25209	0.24235
		$\hat{\alpha}$	0.22722	0.23583	0.23880	0.24617	0.23695
		$\hat{\theta}$	0.28527	0.23646	0.24264	0.21960	0.22995
	AEs	$\hat{\lambda}$	3.54434	3.48021	3.48944	3.47791	3.48086
		$\hat{\beta}$	0.86189	0.78799	0.80407	0.77495	0.77411
		$\hat{\alpha}$	1.46098	1.55211	1.56165	1.51426	1.54771
		$\hat{\theta}$	2.98947	2.70341	2.74405	2.63074	2.64750
	ABBs	$\hat{\lambda}$	0.47709	0.46473	0.46416	0.46200	0.47037
		$\hat{\beta}$	0.19286	0.17555	0.18317	0.18465	0.17937
		$\hat{\alpha}$	0.31610	0.33989	0.34155	0.34946	0.34809
		$\hat{\theta}$	0.68274	0.55341	0.53421	0.52160	0.55095
	MSEs	$\hat{\lambda}$	0.23450	0.22657	0.22679	0.22499	0.23008
		$\hat{\beta}$	0.04354	0.03776	0.04047	0.04050	0.03890
		$\hat{\alpha}$	0.12995	0.14295	0.14620	0.15263	0.14961
		$\hat{\theta}$	0.78596	0.49393	0.47391	0.42686	0.48991

Table 3 (continued)

n	Est.	Est. Par.	MLE	ADE	CVME	LSE	WLSE
50	MREs	$\hat{\lambda}$	0.13631	0.13278	0.13262	0.13200	0.13182
		$\hat{\beta}$	0.25715	0.23407	0.24422	0.24620	0.23916
		$\hat{\alpha}$	0.21074	0.22659	0.22770	0.23297	0.23206
		$\hat{\theta}$	0.27310	0.22136	0.21368	0.20864	0.22038
	AEs	$\hat{\lambda}$	3.48420	3.44891	3.46880	3.46706	3.42226
		$\hat{\beta}$	0.84124	0.77383	0.79097	0.76563	0.76170
		$\hat{\alpha}$	1.46864	1.56256	1.55456	1.53940	1.57001
		$\hat{\theta}$	2.93522	2.66253	2.70476	2.60559	2.62870
	ABBs	$\hat{\lambda}$	0.45557	0.46138	0.45818	0.45881	0.46743
		$\hat{\beta}$	0.18050	0.16962	0.17687	0.17914	0.17381
		$\hat{\alpha}$	0.30481	0.33234	0.33565	0.33981	0.33697
		$\hat{\theta}$	0.65419	0.52981	0.52216	0.51142	0.53310
	MSEs	$\hat{\lambda}$	0.22124	0.22465	0.22254	0.22223	0.22787
		$\hat{\beta}$	0.03932	0.03575	0.03798	0.03814	0.03644
		$\hat{\alpha}$	0.12040	0.13672	0.13851	0.14234	0.13880
		$\hat{\theta}$	0.68648	0.44853	0.44195	0.40228	0.44850
100	MREs	$\hat{\lambda}$	0.13016	0.13182	0.13091	0.13109	0.13355
		$\hat{\beta}$	0.24067	0.22616	0.23583	0.23885	0.23175
		$\hat{\alpha}$	0.20320	0.22156	0.22377	0.22654	0.22465
		$\hat{\theta}$	0.26168	0.21193	0.20887	0.20457	0.21324
	AEs	$\hat{\lambda}$	3.46434	3.49114	3.49617	3.48573	3.48138
		$\hat{\beta}$	0.82486	0.77743	0.79404	0.76735	0.77365
		$\hat{\alpha}$	1.46594	1.53252	1.52438	1.53865	1.53659
		$\hat{\theta}$	2.86677	2.63901	2.69151	2.59591	2.63248
	ABBs	$\hat{\lambda}$	0.43473	0.45539	0.45718	0.44898	0.46036
		$\hat{\beta}$	0.16379	0.15676	0.17323	0.17625	0.16600
		$\hat{\alpha}$	0.28156	0.29216	0.31753	0.32566	0.30546
		$\hat{\theta}$	0.61248	0.49417	0.52033	0.50297	0.52196
	MSEs	$\hat{\lambda}$	0.20712	0.22107	0.22125	0.21641	0.22330
		$\hat{\beta}$	0.03389	0.03136	0.03530	0.03653	0.03393
		$\hat{\alpha}$	0.10289	0.11068	0.12231	0.12846	0.11785
		$\hat{\theta}$	0.59312	0.38493	0.43046	0.37688	0.41540
200	MREs	$\hat{\lambda}$	0.12421	0.13011	0.13059	0.12828	0.13153
		$\hat{\beta}$	0.21839	0.20902	0.22931	0.23499	0.22134
		$\hat{\alpha}$	0.18770	0.19477	0.21169	0.21711	0.20364
		$\hat{\theta}$	0.24499	0.19767	0.19155	0.20119	0.20878
	AEs	$\hat{\lambda}$	3.45888	3.48395	3.49535	3.49008	3.48466
		$\hat{\beta}$	0.80864	0.77811	0.79329	0.76738	0.77812
		$\hat{\alpha}$	1.46850	1.51662	1.50986	1.53413	1.51692
		$\hat{\theta}$	2.81096	2.66719	2.70223	2.60960	2.66916
	ABBs	$\hat{\lambda}$	0.43699	0.44175	0.45634	0.44513	0.43977
		$\hat{\beta}$	0.13988	0.14369	0.17044	0.16312	0.15119
		$\hat{\alpha}$	0.24322	0.26079	0.30074	0.29726	0.27266
		$\hat{\theta}$	0.56029	0.51587	0.49363	0.49213	0.51789
	MSEs	$\hat{\lambda}$	0.20775	0.21188	0.22060	0.21323	0.21102
		$\hat{\beta}$	0.02645	0.02719	0.03494	0.03257	0.02912
		$\hat{\alpha}$	0.07918	0.09152	0.11117	0.11094	0.09674
		$\hat{\theta}$	0.49958	0.41744	0.42426	0.35570	0.40321
400	MREs	$\hat{\lambda}$	0.12485	0.12621	0.13038	0.12718	0.12565
		$\hat{\beta}$	0.18651	0.19158	0.22725	0.21750	0.20159
		$\hat{\alpha}$	0.16215	0.17386	0.20049	0.19817	0.18177
		$\hat{\theta}$	0.22412	0.20635	0.17369	0.19485	0.20016
	AEs	$\hat{\lambda}$	3.40977	3.49330	3.49301	3.49528	3.49212
		$\hat{\beta}$	0.79701	0.77285	0.78897	0.77691	0.77740
		$\hat{\alpha}$	1.47092	1.50701	1.50078	1.50993	1.50195
		$\hat{\theta}$	2.79809	2.64219	2.69488	2.65183	2.66125
	ABBs	$\hat{\lambda}$	0.40179	0.43012	0.43453	0.43481	0.41370
		$\hat{\beta}$	0.11238	0.11757	0.15228	0.14799	0.12648
		$\hat{\alpha}$	0.18542	0.21003	0.25919	0.25921	0.22225
		$\hat{\theta}$	0.51760	0.46092	0.44598	0.46982	0.48987
	MSEs	$\hat{\lambda}$	0.18786	0.20605	0.20719	0.20673	0.20246
		$\hat{\beta}$	0.01912	0.02001	0.02972	0.02824	0.02232
		$\hat{\alpha}$	0.05192	0.06431	0.08804	0.08823	0.06930
		$\hat{\theta}$	0.45702	0.34780	0.41094	0.31580	0.37366
	MREs	$\hat{\lambda}$	0.11480	0.12289	0.12415	0.12423	0.10677
		$\hat{\beta}$	0.14984	0.15676	0.20304	0.19732	0.16864
		$\hat{\alpha}$	0.12361	0.14002	0.17279	0.17281	0.14817
		$\hat{\theta}$	0.20704	0.18437	0.13065	0.16590	0.19595

Table 4Simulation results of the AEs, ABBs, MSEs and MREs of the OLLLW distribution for ($\lambda = 1.5$, $\beta = 1.75$, $\alpha = 1.5$, $\theta = 2.5$).

n	Est.	Est. Par.	MLE	ADE	CVME	LSE	WLSE
20	AEs	$\hat{\lambda}$	1.67118	1.62893	1.62859	1.60130	1.63080
		$\hat{\beta}$	1.82043	1.67441	1.73593	1.65012	1.62134
		$\hat{\alpha}$	1.52594	1.60519	1.60013	1.57086	1.59288
		$\hat{\theta}$	2.25130	2.21103	2.24074	2.20566	2.15454
	ABBs	$\hat{\lambda}$	0.34568	0.38125	0.37960	0.40829	0.40482
		$\hat{\beta}$	0.26676	0.29186	0.27827	0.31683	0.32190
		$\hat{\alpha}$	0.33449	0.36924	0.37299	0.39652	0.38300
		$\hat{\theta}$	0.65242	0.72251	0.82364	0.67691	0.75381
	MSEs	$\hat{\lambda}$	0.14759	0.17231	0.17315	0.19398	0.18986
		$\hat{\beta}$	0.09753	0.11846	0.10261	0.14092	0.14742
		$\hat{\alpha}$	0.14325	0.16749	0.17093	0.18795	0.17735
		$\hat{\theta}$	0.52459	0.63664	0.55614	0.60349	0.69667
	MREs	$\hat{\lambda}$	0.25698	0.25417	0.25307	0.27219	0.26988
		$\hat{\beta}$	0.16823	0.16678	0.15901	0.18105	0.18394
		$\hat{\alpha}$	0.22299	0.24616	0.24866	0.26435	0.25533
		$\hat{\theta}$	0.26097	0.28900	0.46243	0.27076	0.40152
30	AEs	$\hat{\lambda}$	1.65463	1.60902	1.59409	1.58006	1.60762
		$\hat{\beta}$	1.78781	1.65221	1.70467	1.64036	1.62231
		$\hat{\alpha}$	1.52835	1.61895	1.61149	1.59412	1.61250
		$\hat{\theta}$	2.26799	2.24242	2.29199	2.25364	2.21051
	ABBs	$\hat{\lambda}$	0.33568	0.36597	0.36502	0.39227	0.38737
		$\hat{\beta}$	0.25254	0.28739	0.27746	0.30743	0.31532
		$\hat{\alpha}$	0.30434	0.35184	0.35367	0.36822	0.36484
		$\hat{\theta}$	0.64430	0.70017	0.79064	0.66879	0.72541
	MSEs	$\hat{\lambda}$	0.14009	0.16141	0.16207	0.18007	0.17463
		$\hat{\beta}$	0.92364	0.11598	0.10118	0.12767	0.13510
		$\hat{\alpha}$	0.12293	0.15173	0.10036	0.16554	0.16126
		$\hat{\theta}$	0.51302	0.62742	0.42523	0.57262	0.64078
	MREs	$\hat{\lambda}$	0.22378	0.24398	0.24335	0.26151	0.25825
		$\hat{\beta}$	0.14431	0.16799	0.15855	0.17568	0.18018
		$\hat{\alpha}$	0.20290	0.23456	0.23578	0.24548	0.24322
		$\hat{\theta}$	0.25772	0.28007	0.40243	0.26752	0.38103
50	AEs	$\hat{\lambda}$	1.62621	1.59909	1.58694	1.58518	1.59894
		$\hat{\beta}$	1.75500	1.62817	1.67240	1.62225	1.60466
		$\hat{\alpha}$	1.53790	1.62917	1.61117	1.60751	1.63321
		$\hat{\theta}$	2.30873	2.24960	2.29271	2.25148	2.23593
	ABBs	$\hat{\lambda}$	0.30329	0.35146	0.35242	0.37364	0.36282
		$\hat{\beta}$	0.25093	0.27385	0.27740	0.30245	0.31140
		$\hat{\alpha}$	0.29121	0.33748	0.33737	0.34812	0.34765
		$\hat{\theta}$	0.63464	0.69767	0.73956	0.57649	0.71817
	MSEs	$\hat{\lambda}$	0.12187	0.14992	0.15420	0.16805	0.15756
		$\hat{\beta}$	0.08031	0.11681	0.10319	0.12454	0.13453
		$\hat{\alpha}$	0.11209	0.14306	0.14268	0.14960	0.14875
		$\hat{\theta}$	0.50510	0.61742	0.40692	0.49496	0.63704
	MREs	$\hat{\lambda}$	0.20220	0.23431	0.23495	0.24910	0.24188
		$\hat{\beta}$	0.14339	0.16792	0.15851	0.17283	0.17795
		$\hat{\alpha}$	0.19414	0.22499	0.22491	0.23208	0.23177
		$\hat{\theta}$	0.25386	0.27907	0.31684	0.22085	0.32048
1000	AEs	$\hat{\lambda}$	1.62533	1.62120	1.60816	1.60981	1.62067
		$\hat{\beta}$	1.72280	1.62042	1.66680	1.63662	1.61695
		$\hat{\alpha}$	1.53573	1.60259	1.58090	1.58071	1.59907
		$\hat{\theta}$	2.28846	2.19379	2.24384	2.21033	2.19756
	ABBs	$\hat{\lambda}$	0.28104	0.33260	0.35141	0.36136	0.34624
		$\hat{\beta}$	0.24402	0.26905	0.26891	0.27948	0.28079
		$\hat{\alpha}$	0.25451	0.28001	0.30678	0.31008	0.29089
		$\hat{\theta}$	0.63003	0.69036	0.67410	0.48158	0.70870
	MSEs	$\hat{\lambda}$	0.10848	0.13800	0.15023	0.15738	0.14602
		$\hat{\beta}$	0.08089	0.09765	0.09421	0.10306	0.10611
		$\hat{\alpha}$	0.08943	0.10746	0.12028	0.12272	0.11257
		$\hat{\theta}$	0.50036	0.56328	0.37692	0.42096	0.62923
	MREs	$\hat{\lambda}$	0.18736	0.22173	0.23428	0.24091	0.23083
		$\hat{\beta}$	0.13944	0.15374	0.15366	0.15971	0.16045
		$\hat{\alpha}$	0.16967	0.18667	0.20452	0.20672	0.19393
		$\hat{\theta}$	0.25201	0.27640	0.27003	0.20369	0.29148

Table 4 (continued)

n	Est.	Est. Par.	MLE	ADE	CVME	LSE	WLSE
200	AEs	$\hat{\lambda}$	1.60913	1.62931	1.62010	1.62919	1.63196
		$\hat{\beta}$	1.70545	1.62023	1.65700	1.63292	1.62038
		$\hat{\alpha}$	1.53803	1.59251	1.57324	1.57737	1.58788
		$\hat{\theta}$	2.32143	2.19953	2.22888	2.19192	2.19187
	ABBs	$\hat{\lambda}$	0.24243	0.29461	0.32201	0.33374	0.29560
		$\hat{\beta}$	0.23393	0.26805	0.27119	0.28070	0.27109
		$\hat{\alpha}$	0.21064	0.24598	0.27912	0.28214	0.24841
		$\hat{\theta}$	0.60607	0.68798	0.66633	0.38991	0.69639
	MSEs	$\hat{\lambda}$	0.08837	0.11739	0.13175	0.13881	0.11748
		$\hat{\beta}$	0.07736	0.09021	0.09760	0.10567	0.10316
		$\hat{\alpha}$	0.06563	0.08716	0.10323	0.10493	0.08812
		$\hat{\theta}$	0.49680	0.48637	0.24796	0.31065	0.62571
	MREs	$\hat{\lambda}$	0.16162	0.19640	0.21468	0.22249	0.19707
		$\hat{\beta}$	0.13368	0.15317	0.15197	0.16040	0.15491
		$\hat{\alpha}$	0.14043	0.16399	0.18608	0.18809	0.16561
		$\hat{\theta}$	0.24243	0.27519	0.19834	0.17805	0.27856
400	AEs	$\hat{\lambda}$	1.59730	1.62742	1.63220	1.64094	1.62687
		$\hat{\beta}$	1.70274	1.62619	1.64474	1.62583	1.63480
		$\hat{\alpha}$	1.53346	1.57671	1.56757	1.57309	1.56914
		$\hat{\theta}$	2.35961	2.21556	2.21076	2.17330	2.22312
	ABBs	$\hat{\lambda}$	0.20670	0.26024	0.28932	0.29651	0.25943
		$\hat{\beta}$	0.22484	0.24451	0.26355	0.26985	0.24459
		$\hat{\alpha}$	0.16497	0.19331	0.23095	0.23350	0.19642
		$\hat{\theta}$	0.59316	0.67554	0.66595	0.20991	0.67097
	MSEs	$\hat{\lambda}$	0.07103	0.09835	0.11517	0.11900	0.09698
		$\hat{\beta}$	0.07321	0.08733	0.09532	0.10047	0.08798
		$\hat{\alpha}$	0.04444	0.05948	0.07736	0.07861	0.06022
		$\hat{\theta}$	0.48572	0.45036	0.20396	0.16903	0.58889
	MREs	$\hat{\lambda}$	0.13780	0.17349	0.19288	0.19767	0.17296
		$\hat{\beta}$	0.12848	0.13972	0.15060	0.15420	0.13977
		$\hat{\alpha}$	0.10998	0.12887	0.15397	0.15567	0.13095
		$\hat{\theta}$	0.23726	0.27022	0.10356	0.13895	0.26839

$$ABBs = \frac{1}{N} \sum_{i=1}^N |\hat{\psi} - \psi|, \quad MSEs = \frac{1}{N} \sum_{i=1}^N (\hat{\psi} - \psi)^2,$$

$$MREs = \frac{1}{N} \sum_{i=1}^N |\hat{\psi} - \psi| / \psi,$$

where $\psi = (\lambda, \beta, \alpha, \theta)'$. the simulation results include the AEs, MSEs, ABBs and MREs for the OLLW parameters were obtained for the five estimation methods and reported in Tables 1–4. It is noted that the estimates of the OLLW parameters obtained using the five estimation methods are quite reliable and close to the true values of the parameters, showing small MSEs, ABBs and MREs in all parameter combinations. The five estimators, MLE, ADE, CVME, LSE and WLSE, show the consistency property, where the MSEs decrease as n increases for all considered parameter combinations. In summary,

we can conclude that the MLE, ADE, CVME, LSE and WLSE approaches perform very good in estimating the parameters the OLLW model.

6. Real Data Applications

In this section, we analyze two real data sets to show the flexibility of the proposed OLLW model. The first data were previously studied by [Lee and Wang \(2003\)](#) and represent the remission times (in months) of 128 bladder cancer patients. The data were studied by [Abouelmagd et al. \(2018\)](#). The second data contain 40 times to failure of turbocharger of one type of engine. The two data sets are listed below.

Remission times of bladder cancer patients data.

0.08	2.09	13.29	0.40	2.26	3.57	5.06	7.09	9.22	13.80	25.74	0.50
3.48	4.87	23.63	0.20	2.23	6.94	8.66	13.11	3.52	4.98	6.97	9.02
3.88	5.32	7.39	10.34	14.83	34.26	0.90	2.69	4.18	5.34	7.59	10.66
2.46	3.64	5.09	7.26	9.47	14.24	25.82	0.51	2.54	3.70	5.17	7.28
15.96	36.66	1.05	2.69	4.23	5.41	7.62	10.75	16.62	43.01	1.19	2.75
9.74	14.76	26.31	0.81	2.62	3.82	5.32	7.32	10.06	14.77	32.15	2.64

(continued on next page)

(continued)

Remission times of bladder cancer patients data.										
11.79		18.10	1.46	4.40	5.85	8.26	11.98	19.13	1.76	3.25
79.05		1.35	2.87	5.62	7.87	11.64	17.36	1.40	3.02	4.34
4.26		5.41	7.63	17.12	46.12	1.26	2.83	4.33	5.49	7.66
21.73		2.07	3.36	6.93	8.37	12.02	2.02	12.07	20.28	2.02
12.03		3.31	4.51	6.54	8.53	8.65	12.63	22.69	3.36	6.76

Time to failure of turbocharger of one type of engine data.										
1.6	2.0	2.6	3.0	8.0	8.1	8.3	8.4	6.7	6.5	
6.0	6.3	4.5	3.9	3.5	5.0	5.1	7.1	5.8	3.9	
4.8	4.6	5.4	5.3	5.6	8.4	8.5	7.3	7.9	6.1	
7.8	6.5	7.0	8.8	7.7	6.0	7.3	7.7	8.7	9	

We compare the OLLW models with some competitive models including Weibull (W), Frechet Weibull (FW) [Teamah et al. \(2020\)](#), transmuted Weibull (TW) [Aryal and Tsokos \(2011\)](#), gamma Weibull (GW) [Provost et al. \(2011\)](#), transmuted exponentiated Weibull (TExW) [Saboor et al. \(2015\)](#), modified Weibull (MW) [Saboor et al. \(2019\)](#) distributions, using some discrimination criteria namely, Akaike information (AKI), Anderson Darling (ANDA), consistent Akaike information (CAKI), Cramér-von Mises (CRVMI) and Kolmogorov-Smirnov (KOSM) with its *p*-Value.

Table 5

Estimated parameters and SEs of the OLLW model and other competing models for cancer and failure time data.

Model		Estimated Parameters (SEs) for cancer data					
OLLLW	$\hat{\alpha} = 3.96443(5.03434)$	$\hat{\beta} = 0.217695(0.761626)$	$\hat{\theta} = 1.32828(8.75778)$	$\hat{\lambda} = 0.014224(0.284016)$			
W	$\hat{\alpha} = 1.04783(0.0675775)$	$\hat{b} = 9.5607(0.852901)$					
FW	$\hat{\alpha} = 1.14459(3.47881)$	$\hat{\beta} = 1.88103(7.60461)$	$\hat{k} = 0.657074(1.99708)$	$\hat{\lambda} = 1.2456(5.62087)$			
TW	$\hat{\alpha} = 1.13331(0.0754979)$	$\hat{b} = 14.6198(2.25455)$	$\hat{\lambda} = 0.744922(0.202307)$				
GW	$\hat{\alpha} = 1.30988(1.49268)$	$\hat{b} = 0.520095(0.195167)$	$\hat{c} = 1.42917(0.841715)$				
TExW	$\hat{\alpha} = 0.744922(0.202475)$	$\hat{b} = 0.0478365(0.0721666)$	$\hat{c} = 1.13331(0.144136)$	$\hat{\lambda} = 1.61416 \times 10^{-10}(0.078443)$			
MW	$\hat{\lambda} = 1.3172 \times 10^{-27}(0.314243)$	$\hat{\beta} = 0.0938878(0.301314)$	$\hat{k} = 0.1.04783(0.117931)$				

Model		Estimated Parameters (SEs) for failure time data					
OLLLW	$\hat{\alpha} = 0.1207915(0.02685001)$	$\hat{\beta} = 5.5590420(4.47936922)$	$\hat{\theta} = 0.5042859(0.29280235)$	$\hat{\lambda} = 2.5157949(5.38025746)$			
W	$\hat{\alpha} = 3.87251(0.517606)$	$\hat{b} = 6.92003(0.294722)$					
FW	$\hat{\alpha} = 0.641217(7.19206)$	$\hat{\beta} = 2.45759(0.56965)$	$\hat{k} = 3.03254(3.40138)$	$\hat{\lambda} = 3.47328(2.03262)$			
TW	$\hat{\alpha} = 3.58789(0.635904)$	$\hat{b} = 6.59412(0.486356)$	$\hat{\lambda} = -0.336759(0.400413)$				
GW	$\hat{\alpha} = 0.000558(0.00325847)$	$\hat{b} = 3.87251(2.43397)$	$\hat{c} = 5.86338 \times 10^{-13}(3.58527)$				
TExW	$\hat{\alpha} = -1.0000(0.41349)$	$\hat{b} = 5.2725 \times 10^{-7}(1.91555 \times 10^{-6})$	$\hat{c} = 7.08952(1.71117)$	$\hat{\lambda} = 0.126024(0.0396921)$			
MW	$\hat{\lambda} = 0.0295991(0.019097)$	$\hat{\beta} = 9.04066 \times 10^{-6}(0.0000264)$	$\hat{k} = 5.78696(1.37153)$				

Table 6

Measures of the OLLW model and other competing models for cancer and failure time data.

Model	AKI	CAKI	ANDA	CRVMI	KOSM	<i>p</i> -Value	Cancer Data	
							Failure Time Data	
OLLLW	826.841	827.166	0.0945263	0.0145323	0.0312801	0.999627		
W	832.174	832.27	0.957709	0.153703	0.0700169	0.556965		
FW	896.002	896.327	6.11825	0.978722	0.140799	0.0125018		
TW	829.917	830.11	0.560038	0.0879162	0.0587652	0.76866		
GW	827.708	827.902	0.299079	0.0450846	0.0468432	0.941492		
TExW	831.917	832.242	0.560038	0.0879163	0.0587652	0.76866		
MW	834.174	834.367	0.957709	0.153703	0.0700169	0.556965		
OLLLW	163.937	165.0799	0.1101666	0.01690258	0.06358482	0.9969685		
W	168.951	169.275	0.658411	0.0814692	0.107703	0.742309		
FW	211.184	212.326	3.52662	0.623063	0.243817	0.0172041		
TW	170.373	171.039	0.587896	0.0691302	0.102282	0.796815		
GW	170.951	171.618	0.658411	0.0814692	0.107703	0.742309		
TExW	167.253	168.395	0.227123	0.0339386	0.0693949	0.990554		
MW	167.961	168.628	0.311013	0.0434882	0.0898908	0.903026		

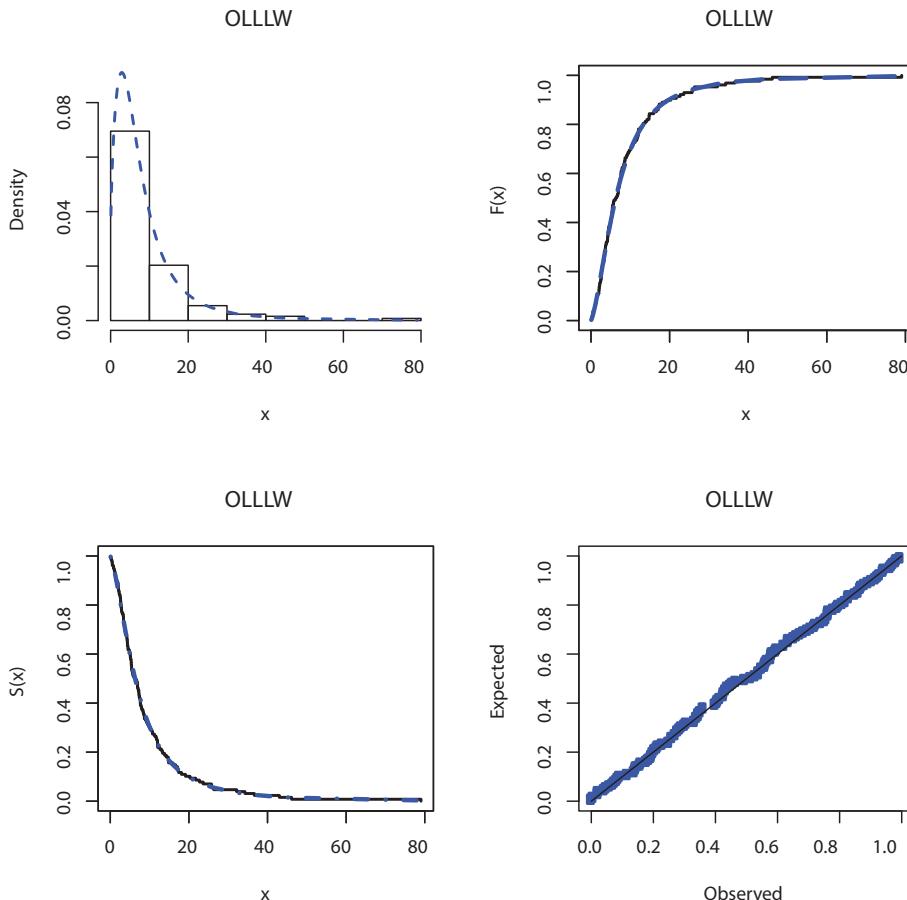


Fig. 3. The fitted PDF, CDF, SF and PP plots of the OLLLW model for cancer data.

Table 5 lists the estimated parameters by the maximum likelihood for both data sets and the standard errors of these estimates for all studied models. The values of the AKI, CAKI, ANDA, CRVMI, KOSM and *p*-Value for the fitted models are shown in **Table 6**. The figures in this table illustrate the superiority of the OLLLW model over other distributions for the two the analyzed sets of data. The plots of fitted PDF, CDF and SF, and probability-probability (PP) plot for the OLLLW distribution are displayed in **Figs. 3 and 4**. These plots support the results in **Table 6** that the new model presents close fit for both data sets.

7. Concluding Remarks

This paper proposed a more flexible extension of the Weibull model called the odd log-logistic Lindley-Weibull (OLLLW) distri-

bution to improve the fitting of the Weibull distribution. The OLLLW density can be expressed as a linear mixture of Weibull densities. The hazard function of the OLLEW provides all important failure rate shapes including bathtub, decreasing, unimodel, increasing, J shaped or reversed-J shaped. Some basic properties are calculated. The parameters of the OLLLW distribution are estimated using some classical estimators including the maximum likelihood, least-squares, Cramér-von Mises, Anderson-Darling and weighted least squares estimators. The simulation results illustrated that the five estimation methods are performing very well for estimating the parameters of the OLLLW model. The importance and flexibility of the OLLLW distribution are studied using two real data sets from medicine and engineering fields. The OLLLW model provides better fit to the analyzed data as compared with other competing models.

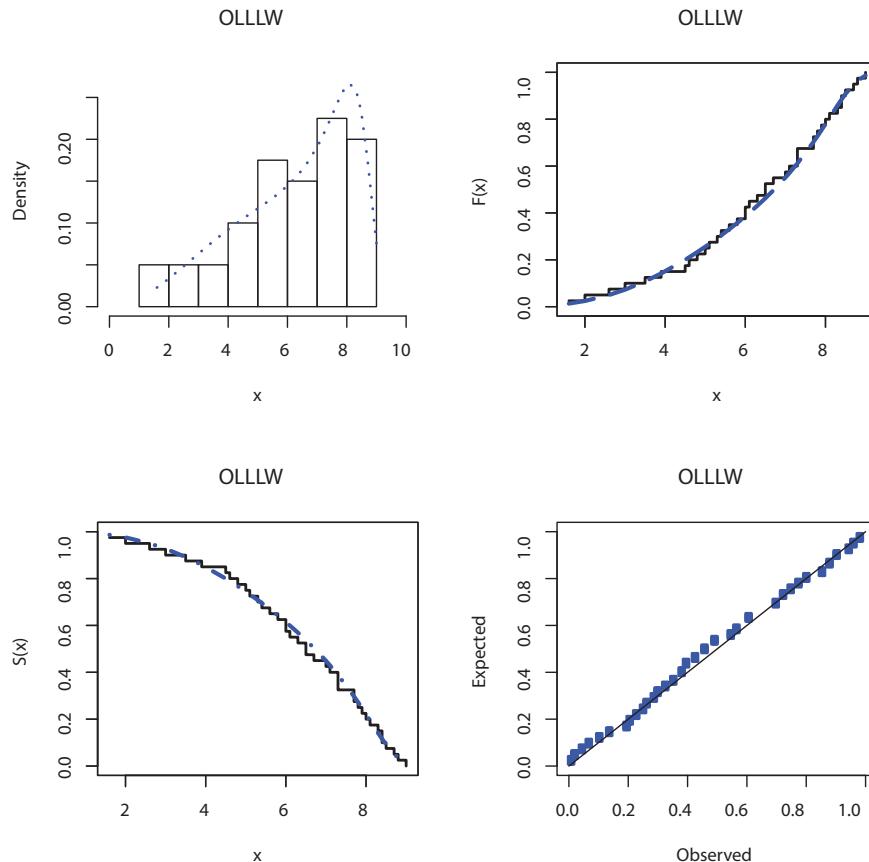


Fig. 4. The fitted PDF, CDF, SF and PP plots of the OLLLW model for failure time data.

Declaration of Competing Interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

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