



Contents lists available at ScienceDirect

Journal of King Saud University – Science

journal homepage: www.sciencedirect.com



Original article

The Marshall–Olkin–Weibull–H family: Estimation, simulations, and applications to COVID-19 data

Ahmed Z. Afify^{a,*}, Hazem Al-Mofleh^b, Hassan M. Aljohani^c, Gauss M. Cordeiro^d^a Department of Statistics, Mathematics and Insurance, Benha University, Benha 13511, Egypt^b Department of Mathematics, Tafila Technical University, Tafila 66110, Jordan^c Department of Mathematics & Statistics, College of Science, Taif University, P.O. Box 11099, Taif 21944, Saudi Arabia^d Department of Statistics, Federal University of Pernambuco, Recife 50710-165, Brazil

ARTICLE INFO

Article history:

Received 3 December 2021

Revised 4 May 2022

Accepted 17 May 2022

Available online 24 May 2022

Keywords:

COVID-19 data

Generalized distribution

Maximum likelihood estimation

Weibull distribution

ABSTRACT

We define a new extended Weibull-H family and obtain some of its mathematical properties. It is very competitive to the beta-G and Kumaraswamy-G classes, which are highly cited in Google Scholar. The parameters of a specified sub-model are estimated by eight methods and its flexibility is proved in two applications to COVID-19 data.

© 2022 The Author(s). Published by Elsevier B.V. on behalf of King Saud University. This is an open access article under the CC BY-NC-ND license (<http://creativecommons.org/licenses/by-nc-nd/4.0/>).

1. Introduction

Adding one or two parameters to parent distributions encourage new concepts for flexible modeling in distribution theory. Among well-established classes of distributions, the exponentiated-G, transmuted-G and Marshall–Olkin-G (MO-G) (Marshall and Olkin, 1997) offer induction of one extra parameter, while the beta-G (Eugene et al., 2002) and Kumaraswamy-G (Cordeiro and de Castro, 2011) classes require two additional shape parameters. Their special cases are explored by Tahir and Nadarajah (2015), among those of other classes.

Composition of distribution generators is emerging as a method to obtain flexible distributions to fit real data in the last five years or so. Some new classes were derived following this method such as the Weibull Marshall–Olkin (Korkmaz et al., 2019), Marshall–Olkin transmuted (Afify et al., 2020), Marshall–Olkin Burr-III (Afify et al., 2021b), among others.

Some other important G families are the exponentiated-G by Gupta et al. (1998), transmuted-G by Shaw and Buckley (2007), gamma-G by Zografos and Balakrishnan (2009), McDonald-G by Alexander et al. (2012), exponentiated-generalized-G by Cordeiro et al. (2013), Burr X-G by Yousof et al. (2017), additive Weibull-G by Hassan and Hemeda (2017), generalized transmuted-G by Nofal et al. (2017), odd Lomax-G by

Cordeiro et al. (2019), Kumaraswamy alpha power-G by Mead et al. (2020), modified Kies-G by Al-Babtain et al. (2020), log-logistic tan-G by Zaidi et al. (2021), and generalized linear failure rate-G by Afify et al. (2022). The interested reader can explore more about parameter induction in Tahir and Nadarajah (2015).

Let $H(x; \xi)$ be a baseline cumulative distribution function (CDF) with a parameter vector ξ . Bourguignon et al. (2014) defined the CDF of the Weibull-H (W-H) class with an extra shape parameter $\beta > 0$ by

$$G(x) = G(x; \beta, \xi) = 1 - \exp \left\{ - \left[\frac{H(x)}{\bar{H}(x)} \right]^\beta \right\}, \quad x \in \mathbb{R}, \quad (1)$$

where $\bar{H}(x) = 1 - H(x)$.

The CDF of the MO-G class is defined by

$$F(x) = F(x; \Omega) = \frac{G(x)}{\theta + (1 - \theta)G(x)}, \quad x \in \mathbb{R}, \quad (2)$$

where $\theta > 0$, $G(x)$ is a parent CDF, and $\Omega = (\theta, \xi^T)^T$.

By combining (1) and (2) (and omitting arguments), the CDF of the Marshall–Olkin–Weibull–H (MOW-H) family (with extra parameters β and θ) follows as

* Corresponding author.

$$F(x) = F(x; \Theta) = \frac{1 - \exp\left\{-\left[\frac{H(x)}{\bar{H}(x)}\right]^\beta\right\}}{1 - (1 - \theta) \exp\left\{-\left[\frac{H(x)}{\bar{H}(x)}\right]^\beta\right\}}, \quad x \in \mathbb{R}, \quad (3)$$

where $\Theta = (\beta, \theta, \xi^T)^T$.

By differentiating (3), the probability density function (PDF) of the MOW-H family reduces to

$$f(x) = f(x; \Theta) = \frac{\theta \beta h(x) H(x)^{\beta-1} \exp\left\{-\left[\frac{H(x)}{\bar{H}(x)}\right]^\beta\right\}}{\bar{H}(x)^{\beta+1} \left\{1 - (1 - \theta) \exp\left\{-\left[\frac{H(x)}{\bar{H}(x)}\right]^\beta\right\}\right\}^2}, \quad (4)$$

where $h(x) = h(x; \xi) = dH(x)/dx$.

Henceforth, $X \sim \text{MOW-H}(\beta, \theta, \xi)$ denotes a random variable (rv) having density (4).

The hazard rate function (HRF) of X is

$$\text{HRF}(x) = \text{HRF}(x; \Theta) = \frac{\beta h(x) H(x)^{\beta-1} \bar{H}(x)^{-\beta-1}}{1 - (1 - \theta) \exp\left\{-\left[\frac{H(x)}{\bar{H}(x)}\right]^\beta\right\}}. \quad (5)$$

By inverting $F(x) = u$ in Eq. (3), we can obtain the quantile function (QF) of X as $x = Q_X(u) = H^{-1}(v/[1 + v])$, where $v = v(u) = [-\log(t)]^{1/\beta}$ and $t = t(u) = (1 - u)/[1 - (1 - \theta)u]$.

A simple interpretation of the MOW-H family can be given as follows. Consider that the variability of the odds $H(x)/\bar{H}(x)$ of a rv Z follows a Weibull distribution with unity scale and shape β . Let N be a positive integer $r\nu$ having a geometric distribution with parameter δ , say $P(N = n) = \delta(1 - \delta)^{n-1}$ (for $n = 1, 2, \dots$). Consider a sequence of N independent copies of Z obtained independently of N . Setting the probability parameters $\delta = \theta$ and $\delta = 1/\theta$ for $\theta \in (0, 1)$ and $\theta > 1$, respectively, the minimum of Z_1, \dots, Z_n has PDF (4).

Furthermore, the proposed MOW-H family extends the Weibull-G class (Bourguignon et al., 2014) which has 536 citations so far, and then it is more flexible than the Weibull-G class. In fact, the plots in Figs. 1 to 8 reveal that the two extra parameters to the baseline model makes the density and risk functions of the new family much more flexible for the four baseline distributions considered here. Additionally, the proposed family can be a competitive generator to the beta-G (Eugene et al., 2002) and Kumaraswamy-G (Cordeiro and de Castro, 2011) classes, which also require two additional shape parameters. These two classes are among the most cited papers in the distribution theory literature.

The paper is structured as follows: Section 2 provides four special cases of Eq. (4), and Section 3 addresses some properties of the new family. Section 4 provides the parameter estimation by eight methods. Two applications to COVID-19 data in Section 6 illustrate the utility of the new family. Section 6 ends with some conclusions.

2. Special Models

This section is devoted to introducing some special sub-models of the MOW-H family. The two extra shape parameters of the MOW-H family make the baseline hazard function more flexible to exhibit all important hazard rate shapes, including monotone and non-monotone shapes.

2.1. MOW-exponential (MOWE)

The MOWE density follows from the exponential $\text{Exp}(a)$ distribution, where $a > 0$. The PDF and CDF of the MOWE distribution are (for $x \in \mathbb{R}$), respectively,

$$f_{\text{MOWE}}(x) = \frac{\theta \beta a [\exp(-ax)]^{-\beta} [1 - \exp(-ax)]^{\beta-1} \exp\{-[\exp(ax) - 1]^\beta\}}{\left\{\theta + (1 - \theta) [1 - \exp\{-[\exp(ax) - 1]^\beta\}]\right\}^2} \quad (6)$$

and

$$F_{\text{MOWE}}(x) = \frac{1 - \exp\{-[\exp(ax) - 1]^\beta\}}{1 - (1 - \theta) \exp\{-[\exp(ax) - 1]^\beta\}}. \quad (7)$$

The QF of the MOWE model (for $0 < u < 1$) reduces to

$$Q_{\text{MOWE}}(u) = -\frac{1}{a} \log \left\{ 1 - \frac{\left(-\log \left[\frac{u-1}{u(1-\theta)-1}\right]\right)^{1/\beta}}{\left(-\log \left[\frac{u-1}{u(1-\theta)-1}\right]\right)^{1/\beta} + 1} \right\}. \quad (8)$$

Fig. 1 displays shapes of the PDF and HRF for some parameters. The HRF can assume increasing, decreasing, reversed-J and J shapes.

2.2. MOW-uniform (MOWU)

For the uniform in the interval $(0, \alpha)$, $H(x) = x/\alpha$, where $\alpha > 0$. Then, the PDF of the MOWU model has the form (for $x \in (0, \alpha)$)

$$f_{\text{MOWU}}(x) = \frac{\theta \beta \alpha^{-1} (x/\alpha)^{\beta-1} \exp\{-[x/(\alpha - x)]^\beta\}}{(1 - x/\alpha)^{\beta+1} \left\{\theta + (1 - \theta) [1 - \exp\{-[x/(\alpha - x)]^\beta\}]\right\}^2}.$$

The Weibull-uniform density when $\theta = 1$ was derived by Phani (1987).

2.3. MOW-Lomax (MOWL)

The Lomax distribution has CDF $H(x) = 1 - (1 + x/\lambda)^{-\alpha}$ (for $x \geq 0$), where $\alpha > 0$ is a shape, and $\lambda > 0$ is a scale. The PDF of the MOWL model (for $x \in \mathbb{R}$) is

$$f_{\text{MOWL}}(x) = \frac{\theta \beta \alpha [1 - (1 + x/\lambda)^{-\alpha}]^{\beta-1} \exp\{-[(1 + x/\lambda)^{\alpha} - 1]^\beta\}}{\lambda (1 + x/\lambda)^{1-\alpha\beta} \left\{\theta + (1 - \theta) [1 - \exp\{-[(1 + x/\lambda)^{\alpha} - 1]^\beta\}]\right\}^2}.$$

2.4. MOW-Weibull (MOWW)

The CDF of the Weibull is $H(x) = 1 - \exp\{-\lambda x^\gamma\}$, where $\lambda > 0$ and $\gamma > 0$, and the MOWW density (for $x \in \mathbb{R}$) has the form

$$f_{\text{MOWW}}(x) = \frac{\theta \beta \lambda \gamma x^{\gamma-1} \exp(-\lambda x^\gamma) [1 - \exp(-\lambda x^\gamma)]^{\beta-1} \exp\{-[\exp(\lambda x^\gamma) - 1]^\beta\}}{\exp[-\lambda(\beta + 1)x^\gamma] \left\{\theta + (1 - \theta) [1 - \exp\{-[\exp(\lambda x^\gamma) - 1]^\beta\}]\right\}^2}. \quad (9)$$

The MOWW model includes the exponential power (Smith and Bain, 1975) and the Chen (2000) distribution when $\theta = \beta = 1$ and $\theta = \beta = \lambda = 1$, respectively.

3. Properties

The PDF associated with the CDF (2) admits the linear representation (Cordeiro et al., 2014)

$$f(x) = \sum_{k=0}^{\infty} w_k \pi_{k+1}(x), \quad (10)$$

where $\pi_{k+1}(x) = (k + 1)g(x)G(x)^k$ is the exponentiated-G (“exp-G”) density with power parameter $k + 1$,

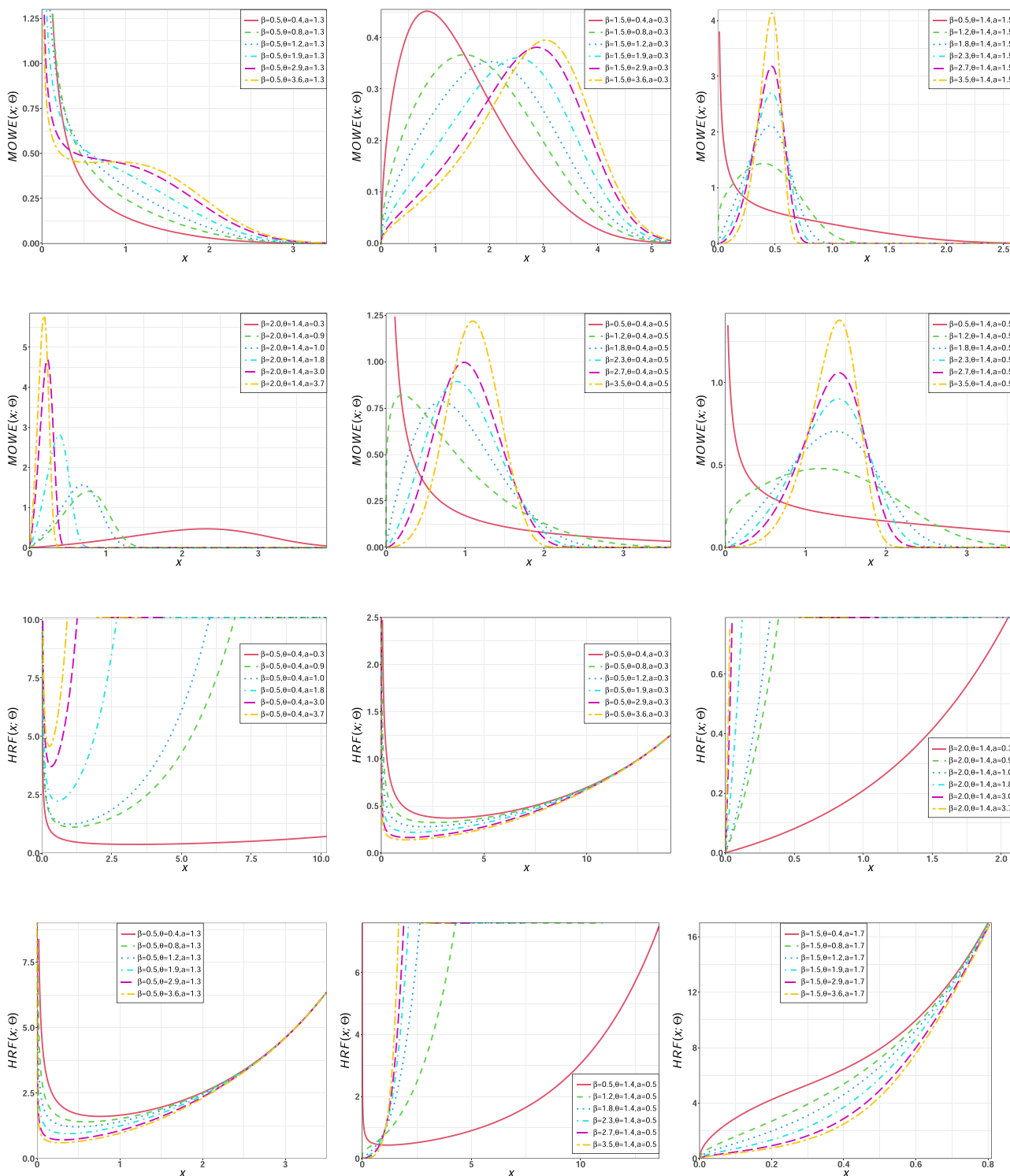


Fig. 1. Shapes of density and hazard functions of the MOWE model for different parameter values.

$$w_k = [p_k \mathbb{1}_{(0,1)}(\theta) + q_k \mathbb{1}_{(1,\infty)}(\theta)],$$

$\mathbb{1}_A(z)$ is the indicator function of a subset A ,

$$p_k = \frac{(-1)^k \theta}{(k+1)} \sum_{j=k}^{\infty} (j+1) \binom{j}{k} \bar{\theta}^j \text{ and } q_k = \theta^{-1} (1 - \theta^{-1})^k.$$

By inserting Eq. (1) and its derivative in $\pi_{k+1}(x)$ and using the expansions for the binomial and exponential function, we obtain (for $k \geq 0$)

$$f(x) = \sum_{r,m=0}^{\infty} q_{r,m} \rho_{(r+1)\beta+m}(x), \tag{11}$$

$$\text{where } \delta_r^{(k)} = \sum_{j=0}^k (-1)^{j+r} (j+1)^r \binom{k}{j},$$

$$q_{r,m} = \beta \sum_{k=0}^{\infty} \frac{(-1)^m (k+1) \binom{-(r+1)\beta-1}{m} \delta_r^{(k)}}{(r+1)\beta+m} w_k,$$

and $\rho_{(r+1)\beta+m}(x)$ is the exp-H density with power $(r + 1)\beta + m$.

So, the density of X is a double linear combination of exp-H densities, which can be adopted with most common type of software, MAPLE, Mathematica, Ox and R, among others.

3.1. Moments

Henceforth, $Z_{(r+1)\beta+m}$ denotes a rv with PDF $\rho_{(r+1)\beta+m}(x)$. Eq. (11) gives

$$\begin{aligned} \mu'_n &= E(X^n) = \sum_{r,m=0}^{\infty} q_{r,m} E[Z_{(r+1)\beta+m}^n] \\ &= \sum_{r,m=0}^{\infty} [(r + 1)\beta + m] q_{r,m} s_{r,m}^{(n)}, \end{aligned} \tag{12}$$

where

$$s_{r,m}^{(n)} = \int_{-\infty}^{\infty} x^n h(x) H(x)^{(r+1)\beta+m-1} dt = \int_0^1 Q_H(u)^n u^{(r+1)\beta+m-1} du$$

can be calculated at least numerically from the exp-H moments or via the baseline QF.

Expressions for several exp-H moments (under special baselines) are reported in many papers such as Nadarajah and Kotz (2006).

The n th incomplete moment of X , say $m_n(y) = E(X|X \leq y)$, follows as

$$m_n(y) = \int_{-\infty}^y x^n f(x) dx = \sum_{r,m=0}^{\infty} [(r + 1)\beta + m] q_{r,m} t_{r,m}^{(n)}(y), \tag{13}$$

where

$$t_{r,m}^{(n)}(y) = \int_{-\infty}^y x^n h(x) H(x)^{(r+1)\beta+m-1} dt = \int_0^{Q_H(y)} Q_H(u)^n u^{(r+1)\beta+m-1} du.$$

The mean deviations and Bonferroni and Lorenz curves of X can be determined from (13) with $n = 1$.

3.2. Generating function

The generating function (gf) of X follows from (11) as

$$\begin{aligned} M(s) &= \sum_{r,m=0}^{\infty} q_{r,m} M_{(r+1)\beta+m}(s) \\ &= \sum_{r,m=0}^{\infty} [(r + 1)\beta + m] [(r + 1)\beta + m] q_{r,m} v_{r,m}(s), \end{aligned}$$

where $M_{(r+1)\beta+m}(s)$ is the gf of the exp-H density $\rho_{(r+1)\beta+m}(x)$

$$v_{r,m}(s) = \int_{-\infty}^{\infty} e^{sx} h(x) H(x)^{(r+1)\beta+m} dx = \int_0^1 \exp\{sQ_H(u)\} u^{(r+1)\beta+m} du.$$

4. Estimation in the MOWE model

Let x_1, \dots, x_n be observations from the MOWE distribution (Section 2.2), and $X_{(1)}, \dots, X_{(n)}$ be the order statistics. The CDF and PDF of this model are denoted by $F(\cdot)$ and $f(\cdot)$, respectively. Its parameters can be estimated by eight methods described below. For more information about these estimation methods, see Nassar et al. (2018), Ramos et al. (2018), Rodrigues et al. (2018), and Ramos et al. (2019).

The ordinary least-squares estimates (OLSEs) minimize the function

$$V(\beta, \theta, a) = \sum_{i=1}^n \left[F(x_i) - \frac{i}{n+1} \right]^2.$$

They can also be found by solving the non-linear equations

$$\sum_{i=1}^n \left[F(x_i) - \frac{i}{n+1} \right] \Delta_s(x_i) = 0, \quad s = 1, 2, 3,$$

where $\Delta_1(x_i) = \Delta_1(x_i|\beta, \theta, a)$, $\Delta_2(x_i) = \Delta_2(x_i|\beta, \theta, a)$ and $\Delta_3(x_i) = \Delta_3(x_i|\beta, \theta, a)$ are

$$\Delta_1(x_i) = \frac{\partial F(x_i)}{\partial \beta} = \theta \frac{(e^{ax_i} - 1)^\beta e^{-(e^{ax_i} - 1)^\beta} \log(e^{ax_i} - 1)}{\left[1 - (1 - \theta)e^{-(e^{ax_i} - 1)^\beta} \right]^2},$$

$$\Delta_2(x_i) = \frac{\partial F(x_i)}{\partial \theta} = - \frac{e^{-(e^{ax_i} - 1)^\beta} \left[1 - e^{-(e^{ax_i} - 1)^\beta} \right]}{\left[1 - (1 - \theta)e^{-(e^{ax_i} - 1)^\beta} \right]^2}$$

and

$$\Delta_3(x_i) = \frac{\partial F(x_i)}{\partial a} = \frac{\beta \theta x_i \left[(e^{ax_i} - 1)^{\beta-1} e^{ax_i - (e^{ax_i} - 1)^\beta} \right]}{\left[1 - (1 - \theta)e^{-(e^{ax_i} - 1)^\beta} \right]^2}.$$

The weighted least-squares estimates (WLSEs) minimize

$$W(\beta, \theta, a) = \sum_{i=1}^n \frac{(n+1)^2(n+2)}{i(n-i+1)} \left[F(x_i) - \frac{i}{n+1} \right]^2,$$

which follow by solving

$$\sum_{i=1}^n \frac{(n+1)^2(n+2)}{i(n-i+1)} \left[F(x_i) - \frac{i}{n+1} \right] \Delta_s(x_i) = 0, \quad s = 1, 2, 3.$$

The maximum likelihood estimates (MLEs) maximize the log-likelihood below for the parameters follows from (6) by classical iterative methods

$$\begin{aligned} \ell(\Theta) &= \sum_{i=1}^n \log(\beta \theta x_i^{\beta-1} + \theta) - \sum_{i=1}^n (\theta x_i^\beta + \theta x_i) \\ &\quad + \sum_{i=1}^n \log(1 + e^{-\theta x_i^\beta - \theta x_i}) - \sum_{i=1}^n (1 - e^{-\theta x_i^\beta - \theta x_i}). \end{aligned} \tag{14}$$

The maximum product of spacing estimates (MPSEs) are good alternatives to the MLEs. Let $D_i = D_i(\beta, \theta, a) = F(x_{(i)}) - F(x_{(i-1)})$ be the uniform spacing (for $i = 1, \dots, n+1$), where $F(x_{(0)}) = 0, F(x_{(n+1)}) = 1$ and $\sum_{i=1}^{n+1} D_i = 1$. The MPSEs maximize the quantity

$$H(\beta, \theta, a) = \sum_{i=1}^{n+1} \log[D_i],$$

which can be determined from the non-linear equations

$$\sum_{i=1}^{n+1} \frac{1}{D_i} [\Delta_s(x_{(i)}) - \Delta_s(x_{(i-1)})] = 0, \quad s = 1, 2, 3.$$

The Cramér-von Mises estimates (CVMs) minimize

$$C(\beta, \theta, a) = - \frac{1}{12n} + \sum_{i=1}^n \left[F(x_i) - \frac{2i-1}{2n} \right]^2,$$

which also follow by solving

$$\sum_{i=1}^n \left[F(x_i) - \frac{2i-1}{2n} \right] \Delta_s(x_i) = 0, \quad s = 1, 2, 3.$$

The Anderson–Darling estimates (ADEs) minimize

$$A(\beta, \theta, a) = -n - \frac{1}{n} \sum_{i=1}^n (2i-1) [\log F(x_i) + \log S(x_i)],$$

which can be found as solutions of the system

$$\sum_{i=1}^n (2i-1) \left[\frac{\Delta_s(x_i)}{F(x_i)} - \frac{\Delta_i(x_{n+1-i})}{S(x_{n+1-i})} \right] = 0, \quad s = 1, 2, 3.$$

The right-tail Anderson–Darling estimates (RADEs) minimize

$$R(\beta, \theta, a) = \frac{n}{2} - 2 \sum_{i=1}^n F(x_{(i)}) - \frac{1}{n} \sum_{i=1}^n (2i-1) \log S(x_{(n+1-i)}).$$

They can also be obtained from the non-linear equations

Table 1
Simulation results for $\beta = 0.2, \theta = 0.1, a = 0.5$.

n	Est.	Est. Par.	WLSE	OLSE	MLE	MPSE	CVME	ADE	RADE	PCE	
30	BIAS	$\hat{\beta}$	0.03046 ⁽³⁾	0.03238 ⁽⁴⁾	0.03296 ⁽⁵⁾	0.01676 ⁽¹⁾	0.03496 ⁽⁶⁾	0.02049 ⁽²⁾	0.03627 ⁽⁷⁾	0.13516 ⁽⁸⁾	
		$\hat{\theta}$	0.05229 ⁽⁴⁾	0.05604 ⁽⁶⁾	0.03758 ⁽³⁾	0.03074 ⁽¹⁾	0.05540 ⁽⁵⁾	0.03340 ⁽²⁾	0.05803 ⁽⁷⁾	0.09540 ⁽⁸⁾	
		\hat{a}	0.25949 ⁽⁵⁾	0.24201 ⁽⁴⁾	0.16096 ⁽²⁾	0.11232 ⁽¹⁾	0.26435 ⁽⁶⁾	0.17529 ⁽³⁾	0.42485 ⁽⁷⁾	0.61630 ⁽⁸⁾	
	MSE	$\hat{\beta}$	0.00156 ⁽³⁾	0.00171 ⁽⁴⁾	0.00658 ⁽⁷⁾	0.00073 ⁽¹⁾	0.00215 ⁽⁵⁾	0.00116 ⁽²⁾	0.00227 ⁽⁶⁾	0.03304 ⁽⁸⁾	
		$\hat{\theta}$	0.00503 ⁽³⁾	0.00597 ⁽⁵⁾	0.00658 ⁽⁷⁾	0.00279 ⁽¹⁾	0.00574 ⁽⁴⁾	0.00318 ⁽²⁾	0.00603 ⁽⁶⁾	0.01496 ⁽⁸⁾	
		\hat{a}	0.14750 ⁽⁵⁾	0.14370 ⁽⁴⁾	0.10088 ⁽²⁾	0.05369 ⁽¹⁾	0.17974 ⁽⁶⁾	0.10141 ⁽³⁾	0.28668 ⁽⁷⁾	0.48110 ⁽⁸⁾	
	MRE	$\hat{\beta}$	0.15231 ⁽³⁾	0.16191 ⁽⁴⁾	0.16482 ⁽⁵⁾	0.08382 ⁽¹⁾	0.17479 ⁽⁶⁾	0.10244 ⁽²⁾	0.18136 ⁽⁷⁾	0.67578 ⁽⁸⁾	
		$\hat{\theta}$	0.52285 ⁽⁴⁾	0.56041 ⁽⁶⁾	0.37581 ⁽³⁾	0.30738 ⁽¹⁾	0.55401 ⁽⁵⁾	0.33404 ⁽²⁾	0.58032 ⁽⁷⁾	0.95399 ⁽⁸⁾	
		\hat{a}	0.51897 ⁽⁵⁾	0.48402 ⁽⁴⁾	0.32192 ⁽²⁾	0.22464 ⁽¹⁾	0.52870 ⁽⁶⁾	0.35058 ⁽³⁾	0.84970 ⁽⁷⁾	1.23259 ⁽⁸⁾	
	\sum Ranks		35 ⁽³⁾	41 ⁽⁵⁾	36 ⁽⁴⁾	9 ⁽¹⁾	49 ⁽⁶⁾	21 ⁽²⁾	61 ⁽⁷⁾	72 ⁽⁸⁾	
	50	BIAS	$\hat{\beta}$	0.02321 ⁽⁴⁾	0.02559 ⁽⁵⁾	0.02135 ⁽³⁾	0.00978 ⁽¹⁾	0.02579 ⁽⁶⁾	0.01157 ⁽²⁾	0.02691 ⁽⁷⁾	0.13480 ⁽⁸⁾
			$\hat{\theta}$	0.04096 ⁽⁴⁾	0.04494 ⁽⁷⁾	0.02332 ⁽³⁾	0.01626 ⁽¹⁾	0.04267 ⁽⁵⁾	0.01885 ⁽²⁾	0.04381 ⁽⁶⁾	0.07792 ⁽⁸⁾
\hat{a}			0.22675 ⁽⁵⁾	0.22534 ⁽⁴⁾	0.07521 ⁽²⁾	0.04893 ⁽¹⁾	0.23677 ⁽⁶⁾	0.10847 ⁽³⁾	0.34806 ⁽⁷⁾	0.61778 ⁽⁸⁾	
MSE		$\hat{\beta}$	0.00088 ⁽³⁾	0.00106 ⁽⁴⁾	0.00460 ⁽⁷⁾	0.00034 ⁽¹⁾	0.00111 ⁽⁵⁾	0.00046 ⁽²⁾	0.00121 ⁽⁶⁾	0.03037 ⁽⁸⁾	
		$\hat{\theta}$	0.00289 ⁽³⁾	0.00369 ⁽⁶⁾	0.00502 ⁽⁷⁾	0.00092 ⁽¹⁾	0.00317 ⁽⁴⁾	0.00123 ⁽²⁾	0.00330 ⁽⁵⁾	0.00934 ⁽⁸⁾	
		\hat{a}	0.12574 ⁽⁴⁾	0.13727 ⁽⁵⁾	0.03648 ⁽²⁾	0.01628 ⁽¹⁾	0.15866 ⁽⁶⁾	0.05574 ⁽³⁾	0.21190 ⁽⁷⁾	0.48144 ⁽⁸⁾	
MRE		$\hat{\beta}$	0.11607 ⁽⁴⁾	0.12793 ⁽⁵⁾	0.10674 ⁽³⁾	0.04888 ⁽¹⁾	0.12894 ⁽⁶⁾	0.05786 ⁽²⁾	0.13456 ⁽⁷⁾	0.67399 ⁽⁸⁾	
		$\hat{\theta}$	0.40965 ⁽⁴⁾	0.44943 ⁽⁷⁾	0.23319 ⁽³⁾	0.16259 ⁽¹⁾	0.42672 ⁽⁵⁾	0.18848 ⁽²⁾	0.43811 ⁽⁶⁾	0.77924 ⁽⁸⁾	
		\hat{a}	0.45350 ⁽⁵⁾	0.45068 ⁽⁴⁾	0.15041 ⁽²⁾	0.09786 ⁽¹⁾	0.47353 ⁽⁶⁾	0.21694 ⁽³⁾	0.69612 ⁽⁷⁾	1.23557 ⁽⁸⁾	
\sum Ranks			36 ⁽⁴⁾	47 ⁽⁵⁾	32 ⁽³⁾	9 ⁽¹⁾	49 ⁽⁶⁾	21 ⁽²⁾	58 ⁽⁷⁾	72 ⁽⁸⁾	
80		BIAS	$\hat{\beta}$	0.01809 ⁽⁴⁾	0.01970 ⁽⁵⁾	0.01470 ⁽³⁾	0.00509 ⁽¹⁾	0.02059 ⁽⁶⁾	0.00615 ⁽²⁾	0.02141 ⁽⁷⁾	0.13888 ⁽⁸⁾
			$\hat{\theta}$	0.03232 ⁽⁴⁾	0.03435 ⁽⁵⁾	0.01345 ⁽³⁾	0.00809 ⁽¹⁾	0.03504 ⁽⁷⁾	0.00991 ⁽²⁾	0.03492 ⁽⁶⁾	0.06194 ⁽⁸⁾
	\hat{a}		0.20932 ⁽⁵⁾	0.19572 ⁽⁴⁾	0.03339 ⁽²⁾	0.01948 ⁽¹⁾	0.21641 ⁽⁶⁾	0.06018 ⁽³⁾	0.26982 ⁽⁷⁾	0.54785 ⁽⁸⁾	
	MSE	$\hat{\beta}$	0.00053 ⁽³⁾	0.00062 ⁽⁴⁾	0.00445 ⁽⁷⁾	0.00014 ⁽¹⁾	0.00069 ⁽⁵⁾	0.00020 ⁽²⁾	0.00075 ⁽⁶⁾	0.03267 ⁽⁸⁾	
		$\hat{\theta}$	0.00177 ⁽³⁾	0.00203 ^(4,5)	0.00280 ⁽⁷⁾	0.00034 ⁽¹⁾	0.00204 ⁽⁶⁾	0.00050 ⁽²⁾	0.00203 ^(4,5)	0.00594 ⁽⁸⁾	
		\hat{a}	0.11569 ⁽⁴⁾	0.11685 ⁽⁵⁾	0.01272 ⁽²⁾	0.00463 ⁽¹⁾	0.14005 ⁽⁷⁾	0.02705 ⁽³⁾	0.13803 ⁽⁶⁾	0.40737 ⁽⁸⁾	
	MRE	$\hat{\beta}$	0.09043 ⁽⁴⁾	0.09852 ⁽⁵⁾	0.07352 ⁽³⁾	0.02547 ⁽¹⁾	0.10295 ⁽⁶⁾	0.03076 ⁽²⁾	0.10704 ⁽⁷⁾	0.69439 ⁽⁸⁾	
		$\hat{\theta}$	0.32319 ⁽⁴⁾	0.34354 ⁽⁵⁾	0.13450 ⁽³⁾	0.08091 ⁽¹⁾	0.35035 ⁽⁷⁾	0.09908 ⁽²⁾	0.34916 ⁽⁶⁾	0.61944 ⁽⁸⁾	
		\hat{a}	0.41865 ⁽⁵⁾	0.39144 ⁽⁴⁾	0.06677 ⁽²⁾	0.03895 ⁽¹⁾	0.43282 ⁽⁶⁾	0.12036 ⁽³⁾	0.53963 ⁽⁷⁾	1.09570 ⁽⁸⁾	
	\sum Ranks		36 ⁽⁴⁾	41.5 ⁽⁵⁾	32 ⁽³⁾	9 ⁽¹⁾	56 ⁽⁶⁾	21 ⁽²⁾	56.5 ⁽⁷⁾	72 ⁽⁸⁾	
	120	BIAS	$\hat{\beta}$	0.01461 ⁽⁴⁾	0.01619 ⁽⁵⁾	0.00693 ⁽³⁾	0.00216 ⁽¹⁾	0.01654 ⁽⁶⁾	0.00297 ⁽²⁾	0.01736 ⁽⁷⁾	0.13589 ⁽⁸⁾
			$\hat{\theta}$	0.02608 ⁽⁴⁾	0.02947 ⁽⁷⁾	0.00563 ⁽³⁾	0.00324 ⁽¹⁾	0.02919 ⁽⁶⁾	0.00479 ⁽²⁾	0.02868 ⁽⁵⁾	0.05220 ⁽⁸⁾
\hat{a}			0.17090 ⁽⁴⁾	0.18720 ⁽⁵⁾	0.01370 ⁽²⁾	0.00637 ⁽¹⁾	0.19666 ⁽⁶⁾	0.03035 ⁽³⁾	0.21761 ⁽⁷⁾	0.47641 ⁽⁸⁾	
MSE		$\hat{\beta}$	0.00034 ⁽³⁾	0.00042 ⁽⁴⁾	0.00223 ⁽⁷⁾	0.00005 ⁽¹⁾	0.00045 ⁽⁵⁾	0.00008 ⁽²⁾	0.00048 ⁽⁶⁾	0.03406 ⁽⁸⁾	
		$\hat{\theta}$	0.00113 ⁽⁴⁾	0.00149 ⁽⁷⁾	0.00088 ⁽³⁾	0.00011 ⁽¹⁾	0.00141 ⁽⁶⁾	0.00019 ⁽²⁾	0.00132 ⁽⁵⁾	0.00417 ⁽⁸⁾	
		\hat{a}	0.08650 ⁽⁴⁾	0.10942 ⁽⁶⁾	0.00442 ⁽²⁾	0.00091 ⁽¹⁾	0.12331 ⁽⁷⁾	0.01181 ⁽³⁾	0.09265 ⁽⁵⁾	0.33211 ⁽⁸⁾	
MRE		$\hat{\beta}$	0.07304 ⁽⁴⁾	0.08094 ⁽⁵⁾	0.03465 ⁽³⁾	0.01080 ⁽¹⁾	0.08271 ⁽⁶⁾	0.01487 ⁽²⁾	0.08681 ⁽⁷⁾	0.67947 ⁽⁸⁾	
		$\hat{\theta}$	0.26083 ⁽⁴⁾	0.29466 ⁽⁷⁾	0.05631 ⁽³⁾	0.03235 ⁽¹⁾	0.29191 ⁽⁶⁾	0.04786 ⁽²⁾	0.28683 ⁽⁵⁾	0.52202 ⁽⁸⁾	
		\hat{a}	0.34180 ⁽⁴⁾	0.37440 ⁽⁵⁾	0.02740 ⁽²⁾	0.01273 ⁽¹⁾	0.39332 ⁽⁶⁾	0.06069 ⁽³⁾	0.43522 ⁽⁷⁾	0.95282 ⁽⁸⁾	
\sum Ranks			35 ⁽⁴⁾	51 ⁽⁵⁾	28 ⁽³⁾	9 ⁽¹⁾	54 ^(6,5)	21 ⁽²⁾	54 ^(6,5)	72 ⁽⁸⁾	
200		BIAS	$\hat{\beta}$	0.01152 ⁽⁴⁾	0.01277 ⁽⁶⁾	0.00220 ⁽³⁾	0.00045 ⁽¹⁾	0.01256 ⁽⁵⁾	0.00074 ⁽²⁾	0.01306 ⁽⁷⁾	0.12438 ⁽⁸⁾
			$\hat{\theta}$	0.02074 ⁽⁴⁾	0.02385 ⁽⁷⁾	0.00174 ⁽³⁾	0.00065 ⁽¹⁾	0.02336 ⁽⁶⁾	0.00117 ⁽²⁾	0.02155 ⁽⁵⁾	0.04304 ⁽⁸⁾
	\hat{a}		0.13754 ⁽⁴⁾	0.17145 ⁽⁶⁾	0.00326 ⁽²⁾	0.00156 ⁽¹⁾	0.18696 ⁽⁷⁾	0.00731 ⁽³⁾	0.15700 ⁽⁵⁾	0.39447 ⁽⁸⁾	
	MSE	$\hat{\beta}$	0.00021 ⁽³⁾	0.00026 ^(4,5)	0.00083 ⁽⁷⁾	0.00001 ⁽¹⁾	0.00026 ^(4,5)	0.00002 ⁽²⁾	0.00027 ⁽⁶⁾	0.03047 ⁽⁸⁾	
		$\hat{\theta}$	0.00070 ⁽⁴⁾	0.00095 ⁽⁷⁾	0.00045 ⁽³⁾	0.00001 ⁽¹⁾	0.00092 ⁽⁶⁾	0.00004 ⁽²⁾	0.00075 ⁽⁵⁾	0.00279 ⁽⁸⁾	
		\hat{a}	0.05948 ⁽⁵⁾	0.09662 ⁽⁶⁾	0.00107 ⁽²⁾	0.00012 ⁽¹⁾	0.11430 ⁽⁷⁾	0.00225 ⁽³⁾	0.04729 ⁽⁴⁾	0.24448 ⁽⁸⁾	
	MRE	$\hat{\beta}$	0.05760 ⁽⁴⁾	0.06386 ⁽⁶⁾	0.01099 ⁽³⁾	0.00224 ⁽¹⁾	0.06278 ⁽⁵⁾	0.00372 ⁽²⁾	0.06531 ⁽⁷⁾	0.62192 ⁽⁸⁾	
		$\hat{\theta}$	0.20742 ⁽⁴⁾	0.23853 ⁽⁷⁾	0.01741 ⁽³⁾	0.00655 ⁽¹⁾	0.23359 ⁽⁶⁾	0.01174 ⁽²⁾	0.21547 ⁽⁵⁾	0.43040 ⁽⁸⁾	
		\hat{a}	0.27509 ⁽⁴⁾	0.34290 ⁽⁶⁾	0.00651 ⁽²⁾	0.00312 ⁽¹⁾	0.37392 ⁽⁷⁾	0.01462 ⁽³⁾	0.31400 ⁽⁵⁾	0.78894 ⁽⁸⁾	
	\sum Ranks		36 ⁽⁴⁾	55.5 ⁽⁷⁾	28 ⁽³⁾	9 ⁽¹⁾	53.5 ⁽⁶⁾	21 ⁽²⁾	49 ⁽⁵⁾	72 ⁽⁸⁾	

$$- 2 \sum_{i=1}^n \Delta_s(x_{(i)}) + \frac{1}{n} \sum_{i=1}^n (2i - 1) \frac{\Delta_s(x_{(n+1-i)})}{S(x_{(n+1-i)})}$$

Let $u_i = i/(n + 1)$ be an unbiased estimator of $F(x_{(i)})$. The PC estimates (PCEs) minimize

$$P(\beta, \varphi, \rho) = \sum_{i=1}^n \left(x_{(i)} - \frac{1}{\beta} \log \left[\sqrt{-\log \left(1 - \left(\frac{1}{2\rho} [(\rho + 1) - \sqrt{\delta(\rho, u_i)}] \right)^{1/\varphi}} + 1 \right) \right]^2 \right),$$

where $\delta(\rho, u_i) = (\rho + 1)^2 - 4\rho u_i$.

5. Simulation analysis

The simulation study compares the estimates from the eight methods in Section 4 in terms of the averages of the four quantities: absolute bias ($|\text{Bias}(\hat{\theta})|$), $|\text{Bias}(\hat{\theta})| = \frac{1}{N} \sum_{i=1}^N |\hat{\theta} - \theta|$, mean square error (MSE), $\text{MSE} = \frac{1}{N} \sum_{i=1}^N (\hat{\theta} - \theta)^2$, and mean relative error (MRE), $\text{MRE} = \frac{1}{N} \sum_{i=1}^N |\hat{\theta} - \theta|/\theta$.

Table 2
Simulation results for $\beta = 1.0, \theta = 0.1, a = 0.5$.

n	Est.	Est. Par.	WLSE	OLSE	MLE	MPSE	CVME	ADE	RADE	PCE	
30	BIAS	$\hat{\beta}$	0.18821 ⁽⁴⁾	0.19683 ⁽⁵⁾	0.16502 ⁽¹⁾	0.17886 ⁽²⁾	0.20095 ⁽⁶⁾	0.17967 ⁽³⁾	0.20987 ⁽⁷⁾	0.28430 ⁽⁸⁾	
		$\hat{\theta}$	0.14864 ⁽⁵⁾	0.16855 ⁽⁷⁾	0.12346 ⁽¹⁾	0.13625 ⁽³⁾	0.16485 ⁽⁶⁾	0.12721 ⁽²⁾	0.13666 ⁽⁴⁾	0.21117 ⁽⁸⁾	
		\hat{a}	0.38120 ⁽⁵⁾	0.42480 ⁽⁷⁾	0.28647 ⁽¹⁾	0.30892 ⁽²⁾	0.43726 ⁽⁸⁾	0.32772 ⁽⁴⁾	0.32536 ⁽³⁾	0.42004 ⁽⁶⁾	
	MSE	$\hat{\beta}$	0.05515 ⁽⁴⁾	0.06053 ⁽⁵⁾	0.04360 ⁽¹⁾	0.04795 ⁽²⁾	0.06523 ⁽⁶⁾	0.05185 ⁽³⁾	0.07054 ⁽⁷⁾	0.11469 ⁽⁸⁾	
		$\hat{\theta}$	0.05069 ⁽⁵⁾	0.06359 ⁽⁷⁾	0.03664 ⁽¹⁾	0.04405 ⁽³⁾	0.05997 ⁽⁶⁾	0.03776 ⁽²⁾	0.04610 ⁽⁴⁾	0.09969 ⁽⁸⁾	
		\hat{a}	0.21907 ⁽⁵⁾	0.25823 ⁽⁶⁾	0.15469 ⁽¹⁾	0.15513 ⁽²⁾	0.27463 ⁽⁸⁾	0.17103 ⁽⁴⁾	0.17084 ⁽³⁾	0.26387 ⁽⁷⁾	
	MRE	$\hat{\beta}$	0.18821 ⁽⁴⁾	0.19683 ⁽⁵⁾	0.16502 ⁽¹⁾	0.17886 ⁽²⁾	0.20095 ⁽⁶⁾	0.17967 ⁽³⁾	0.20987 ⁽⁷⁾	0.28430 ⁽⁸⁾	
		$\hat{\theta}$	1.48640 ⁽⁵⁾	1.68551 ⁽⁷⁾	1.23455 ⁽¹⁾	1.36251 ⁽³⁾	1.64850 ⁽⁶⁾	1.27207 ⁽²⁾	1.36664 ⁽⁴⁾	2.11173 ⁽⁸⁾	
		\hat{a}	0.76240 ⁽⁵⁾	0.84960 ⁽⁷⁾	0.57293 ⁽¹⁾	0.61784 ⁽²⁾	0.87451 ⁽⁸⁾	0.65544 ⁽⁴⁾	0.65072 ⁽³⁾	0.84008 ⁽⁶⁾	
	\sum Ranks		42 ^(4.5)	56 ⁽⁶⁾	9 ⁽¹⁾	21 ⁽²⁾	60 ⁽⁷⁾	27 ⁽³⁾	42 ^(4.5)	67 ⁽⁸⁾	
	50	BIAS	$\hat{\beta}$	0.14676 ⁽⁴⁾	0.15674 ⁽⁵⁾	0.13258 ⁽¹⁾	0.13664 ⁽²⁾	0.16032 ⁽⁶⁾	0.14150 ⁽³⁾	0.16366 ⁽⁷⁾	0.26690 ⁽⁸⁾
			$\hat{\theta}$	0.11156 ⁽⁵⁾	0.13187 ⁽⁷⁾	0.09356 ⁽³⁾	0.09311 ⁽²⁾	0.12731 ⁽⁶⁾	0.09185 ⁽¹⁾	0.09686 ⁽⁴⁾	0.15813 ⁽⁸⁾
\hat{a}			0.29760 ⁽⁵⁾	0.35585 ⁽⁸⁾	0.21439 ⁽¹⁾	0.21660 ⁽²⁾	0.35580 ⁽⁷⁾	0.24653 ⁽⁴⁾	0.23950 ⁽³⁾	0.33033 ⁽⁶⁾	
MSE		$\hat{\beta}$	0.03346 ⁽⁴⁾	0.03821 ⁽⁵⁾	0.02771 ⁽¹⁾	0.02883 ⁽²⁾	0.04077 ⁽⁶⁾	0.03197 ⁽³⁾	0.04317 ⁽⁷⁾	0.10387 ⁽⁸⁾	
		$\hat{\theta}$	0.02810 ⁽⁵⁾	0.03761 ⁽⁷⁾	0.02082 ⁽³⁾	0.02072 ⁽²⁾	0.03516 ⁽⁶⁾	0.01886 ⁽¹⁾	0.02242 ⁽⁴⁾	0.05794 ⁽⁸⁾	
		\hat{a}	0.14532 ⁽⁵⁾	0.19325 ⁽⁷⁾	0.08972 ⁽²⁾	0.08137 ⁽¹⁾	0.19647 ⁽⁸⁾	0.10260 ⁽⁴⁾	0.09795 ⁽³⁾	0.17628 ⁽⁶⁾	
MRE		$\hat{\beta}$	0.14676 ⁽⁴⁾	0.15674 ⁽⁵⁾	0.13258 ⁽¹⁾	0.13664 ⁽²⁾	0.16032 ⁽⁶⁾	0.14150 ⁽³⁾	0.16366 ⁽⁷⁾	0.26690 ⁽⁸⁾	
		$\hat{\theta}$	1.11564 ⁽⁵⁾	1.31869 ⁽⁷⁾	0.93559 ⁽³⁾	0.93114 ⁽²⁾	1.27312 ⁽⁶⁾	0.91853 ⁽¹⁾	0.96858 ⁽⁴⁾	1.58129 ⁽⁸⁾	
		\hat{a}	0.59521 ⁽⁵⁾	0.71169 ⁽⁸⁾	0.42878 ⁽¹⁾	0.43319 ⁽²⁾	0.71161 ⁽⁷⁾	0.49306 ⁽⁴⁾	0.47899 ⁽³⁾	0.66066 ⁽⁶⁾	
\sum Ranks			42 ^(4.5)	59 ⁽⁷⁾	16 ⁽¹⁾	17 ⁽²⁾	58 ⁽⁶⁾	24 ⁽³⁾	42 ^(4.5)	66 ⁽⁸⁾	
80		BIAS	$\hat{\beta}$	0.11814 ⁽⁴⁾	0.13003 ⁽⁵⁾	0.10279 ⁽¹⁾	0.10867 ⁽²⁾	0.13088 ⁽⁶⁾	0.11319 ⁽³⁾	0.13094 ⁽⁷⁾	0.23106 ⁽⁸⁾
			$\hat{\theta}$	0.08138 ⁽⁵⁾	0.10604 ⁽⁷⁾	0.06503 ⁽¹⁾	0.06655 ⁽²⁾	0.10259 ⁽⁶⁾	0.07122 ⁽³⁾	0.07425 ⁽⁴⁾	0.11504 ⁽⁸⁾
	\hat{a}		0.21924 ⁽⁵⁾	0.29265 ⁽⁸⁾	0.15519 ⁽¹⁾	0.15867 ⁽²⁾	0.28978 ⁽⁷⁾	0.18915 ⁽⁴⁾	0.18405 ⁽³⁾	0.24647 ⁽⁶⁾	
	MSE	$\hat{\beta}$	0.02205 ⁽⁴⁾	0.02636 ⁽⁵⁾	0.01657 ⁽¹⁾	0.01850 ⁽²⁾	0.02692 ⁽⁶⁾	0.02065 ⁽³⁾	0.02779 ⁽⁷⁾	0.08225 ⁽⁸⁾	
		$\hat{\theta}$	0.01448 ⁽⁵⁾	0.02506 ⁽⁷⁾	0.00920 ⁽¹⁾	0.00943 ⁽²⁾	0.02263 ⁽⁶⁾	0.01061 ⁽³⁾	0.01231 ⁽⁴⁾	0.03207 ⁽⁸⁾	
		\hat{a}	0.08309 ⁽⁵⁾	0.13868 ⁽⁸⁾	0.04647 ⁽²⁾	0.04259 ⁽¹⁾	0.13567 ⁽⁷⁾	0.06138 ⁽⁴⁾	0.05862 ⁽³⁾	0.10632 ⁽⁶⁾	
	MRE	$\hat{\beta}$	0.11814 ⁽⁴⁾	0.13003 ⁽⁵⁾	0.10279 ⁽¹⁾	0.10867 ⁽²⁾	0.13088 ⁽⁶⁾	0.11319 ⁽³⁾	0.13094 ⁽⁷⁾	0.23106 ⁽⁸⁾	
		$\hat{\theta}$	0.81384 ⁽⁵⁾	1.06043 ⁽⁷⁾	0.65030 ⁽¹⁾	0.66551 ⁽²⁾	1.02589 ⁽⁶⁾	0.71220 ⁽³⁾	0.74247 ⁽⁴⁾	1.15041 ⁽⁸⁾	
		\hat{a}	0.43848 ⁽⁵⁾	0.58531 ⁽⁸⁾	0.31038 ⁽¹⁾	0.31734 ⁽²⁾	0.57957 ⁽⁷⁾	0.37830 ⁽⁴⁾	0.36809 ⁽³⁾	0.49293 ⁽⁶⁾	
	\sum Ranks		42 ^(4.5)	60 ⁽⁷⁾	10 ⁽¹⁾	17 ⁽²⁾	57 ⁽⁶⁾	30 ⁽³⁾	42 ^(4.5)	66 ⁽⁸⁾	
	120	BIAS	$\hat{\beta}$	0.09436 ⁽⁴⁾	0.10586 ⁽⁶⁾	0.08534 ⁽¹⁾	0.08578 ⁽²⁾	0.10595 ⁽⁷⁾	0.09157 ⁽³⁾	0.10412 ⁽⁵⁾	0.20363 ⁽⁸⁾
			$\hat{\theta}$	0.06303 ⁽⁵⁾	0.08109 ⁽⁶⁾	0.05006 ⁽²⁾	0.04924 ⁽¹⁾	0.08173 ⁽⁷⁾	0.05536 ⁽³⁾	0.05740 ⁽⁴⁾	0.09080 ⁽⁸⁾
\hat{a}			0.16761 ⁽⁵⁾	0.23535 ⁽⁷⁾	0.11845 ⁽¹⁾	0.12295 ⁽²⁾	0.23821 ⁽⁸⁾	0.15035 ⁽⁴⁾	0.14345 ⁽³⁾	0.19543 ⁽⁶⁾	
MSE		$\hat{\beta}$	0.01398 ⁽⁴⁾	0.01746 ⁽⁶⁾	0.01147 ⁽¹⁾	0.01148 ⁽²⁾	0.01765 ⁽⁷⁾	0.01325 ⁽³⁾	0.01745 ⁽⁵⁾	0.06719 ⁽⁸⁾	
		$\hat{\theta}$	0.00809 ⁽⁵⁾	0.01404 ⁽⁷⁾	0.00512 ⁽²⁾	0.00470 ⁽¹⁾	0.01374 ⁽⁶⁾	0.00591 ⁽³⁾	0.00675 ⁽⁴⁾	0.01973 ⁽⁸⁾	
		\hat{a}	0.05056 ⁽⁵⁾	0.09200 ⁽⁷⁾	0.02575 ⁽²⁾	0.02490 ⁽¹⁾	0.09515 ⁽⁸⁾	0.03856 ⁽⁴⁾	0.03566 ⁽³⁾	0.06859 ⁽⁶⁾	
MRE		$\hat{\beta}$	0.09436 ⁽⁴⁾	0.10586 ⁽⁶⁾	0.08534 ⁽¹⁾	0.08578 ⁽²⁾	0.10595 ⁽⁷⁾	0.09157 ⁽³⁾	0.10412 ⁽⁵⁾	0.20363 ⁽⁸⁾	
		$\hat{\theta}$	0.63028 ⁽⁵⁾	0.81089 ⁽⁶⁾	0.50063 ⁽²⁾	0.49243 ⁽¹⁾	0.81726 ⁽⁷⁾	0.55363 ⁽³⁾	0.57405 ⁽⁴⁾	0.90800 ⁽⁸⁾	
		\hat{a}	0.33523 ⁽⁵⁾	0.47070 ⁽⁷⁾	0.23690 ⁽¹⁾	0.24590 ⁽²⁾	0.47641 ⁽⁸⁾	0.30070 ⁽⁴⁾	0.28691 ⁽³⁾	0.39085 ⁽⁶⁾	
\sum Ranks			42 ⁽⁵⁾	58 ⁽⁶⁾	13 ⁽¹⁾	14 ⁽²⁾	65 ⁽⁷⁾	30 ⁽³⁾	36 ⁽⁴⁾	66 ⁽⁸⁾	
200		BIAS	$\hat{\beta}$	0.07194 ⁽⁴⁾	0.08231 ⁽⁶⁾	0.06480 ⁽¹⁾	0.06632 ⁽²⁾	0.08304 ⁽⁷⁾	0.07017 ⁽³⁾	0.08134 ⁽⁵⁾	0.16421 ⁽⁸⁾
			$\hat{\theta}$	0.04579 ⁽⁵⁾	0.06099 ⁽⁷⁾	0.03568 ⁽¹⁾	0.03624 ⁽²⁾	0.06020 ⁽⁶⁾	0.04150 ⁽³⁾	0.04386 ⁽⁴⁾	0.06734 ⁽⁸⁾
	\hat{a}		0.12452 ⁽⁵⁾	0.17891 ⁽⁸⁾	0.08785 ⁽¹⁾	0.09251 ⁽²⁾	0.17721 ⁽⁷⁾	0.11233 ⁽⁴⁾	0.10904 ⁽³⁾	0.14422 ⁽⁶⁾	
	MSE	$\hat{\beta}$	0.00809 ⁽⁴⁾	0.01076 ⁽⁶⁾	0.00670 ⁽¹⁾	0.00690 ⁽²⁾	0.01083 ⁽⁷⁾	0.00773 ⁽³⁾	0.01040 ⁽⁵⁾	0.04379 ⁽⁸⁾	
		$\hat{\theta}$	0.00399 ⁽⁵⁾	0.00720 ⁽⁷⁾	0.00242 ⁽²⁾	0.00228 ⁽¹⁾	0.00705 ⁽⁶⁾	0.00311 ⁽³⁾	0.00354 ⁽⁴⁾	0.00968 ⁽⁸⁾	
		\hat{a}	0.02621 ⁽⁵⁾	0.05336 ⁽⁸⁾	0.01361 ⁽²⁾	0.01331 ⁽¹⁾	0.05302 ⁽⁷⁾	0.02104 ⁽⁴⁾	0.01999 ⁽³⁾	0.03579 ⁽⁶⁾	
	MRE	$\hat{\beta}$	0.07194 ⁽⁴⁾	0.08231 ⁽⁶⁾	0.06480 ⁽¹⁾	0.06632 ⁽²⁾	0.08304 ⁽⁷⁾	0.07017 ⁽³⁾	0.08134 ⁽⁵⁾	0.16421 ⁽⁸⁾	
		$\hat{\theta}$	0.45795 ⁽⁵⁾	0.60988 ⁽⁷⁾	0.35676 ⁽¹⁾	0.36239 ⁽²⁾	0.60198 ⁽⁶⁾	0.41495 ⁽³⁾	0.43860 ⁽⁴⁾	0.67335 ⁽⁸⁾	
		\hat{a}	0.24904 ⁽⁵⁾	0.35782 ⁽⁸⁾	0.17570 ⁽¹⁾	0.18503 ⁽²⁾	0.35443 ⁽⁷⁾	0.22466 ⁽⁴⁾	0.21809 ⁽³⁾	0.28844 ⁽⁶⁾	
	\sum Ranks		42 ⁽⁵⁾	63 ⁽⁷⁾	11 ⁽¹⁾	16 ⁽²⁾	60 ⁽⁶⁾	30 ⁽³⁾	36 ⁽⁴⁾	66 ⁽⁸⁾	

Table 3
Simulation results for $\beta = 2.1, \theta = 0.1, a = 0.5$.

<i>n</i>	Est.	Est. Par.	WLSE	OLSE	MLE	MPSE	CVME	ADE	RADE	PCE	
30	BIAS	$\hat{\beta}$	0.44714 ⁽⁵⁾	0.46622 ⁽⁸⁾	0.37061 ⁽¹⁾	0.41964 ⁽²⁾	0.45743 ⁽⁶⁾	0.41980 ⁽³⁾	0.46598 ⁽⁷⁾	0.42260 ⁽⁴⁾	
		$\hat{\theta}$	0.23678 ⁽⁶⁾	0.27079 ⁽⁸⁾	0.18572 ⁽¹⁾	0.20457 ⁽⁴⁾	0.27037 ⁽⁷⁾	0.20204 ⁽³⁾	0.21242 ⁽⁵⁾	0.20125 ⁽²⁾	
		\hat{a}	0.24813 ⁽⁶⁾	0.27043 ⁽⁸⁾	0.14955 ⁽¹⁾	0.19904 ⁽²⁾	0.26574 ⁽⁷⁾	0.21153 ⁽⁴⁾	0.20733 ⁽³⁾	0.21483 ⁽⁵⁾	
	MSE	$\hat{\beta}$	0.28217 ⁽⁵⁾	0.30248 ⁽⁷⁾	0.20680 ⁽¹⁾	0.25631 ⁽³⁾	0.29252 ⁽⁶⁾	0.25579 ⁽²⁾	0.30810 ⁽⁸⁾	0.25703 ⁽⁴⁾	
		$\hat{\theta}$	0.14231 ⁽⁶⁾	0.17644 ⁽⁸⁾	0.09604 ⁽¹⁾	0.11278 ⁽⁴⁾	0.17643 ⁽⁷⁾	0.11099 ⁽³⁾	0.12279 ⁽⁵⁾	0.10917 ⁽²⁾	
		\hat{a}	0.08862 ⁽⁶⁾	0.10268 ⁽⁸⁾	0.03955 ⁽¹⁾	0.06237 ⁽²⁾	0.10054 ⁽⁷⁾	0.07012 ⁽⁴⁾	0.06730 ⁽³⁾	0.07119 ⁽⁵⁾	
	MRE	$\hat{\beta}$	0.21292 ⁽⁵⁾	0.22201 ⁽⁸⁾	0.17648 ⁽¹⁾	0.19983 ⁽²⁾	0.21783 ⁽⁶⁾	0.19990 ⁽³⁾	0.22190 ⁽⁷⁾	0.20124 ⁽⁴⁾	
		$\hat{\theta}$	2.36782 ⁽⁶⁾	2.70785 ⁽⁸⁾	1.85723 ⁽¹⁾	2.04574 ⁽⁴⁾	2.70373 ⁽⁷⁾	2.02043 ⁽³⁾	2.12422 ⁽⁵⁾	2.01250 ⁽²⁾	
		\hat{a}	0.49626 ⁽⁶⁾	0.54086 ⁽⁸⁾	0.29910 ⁽¹⁾	0.39808 ⁽²⁾	0.53149 ⁽⁷⁾	0.42305 ⁽⁴⁾	0.41466 ⁽³⁾	0.42967 ⁽⁵⁾	
	$\sum Ranks$			51 ⁽⁶⁾	71 ⁽⁸⁾	9 ⁽¹⁾	25 ⁽²⁾	60 ⁽⁷⁾	29 ⁽³⁾	46 ⁽⁵⁾	33 ⁽⁴⁾
	50	BIAS	$\hat{\beta}$	0.36210 ⁽⁵⁾	0.39106 ⁽⁷⁾	0.29598 ⁽¹⁾	0.32784 ⁽²⁾	0.39737 ⁽⁸⁾	0.34439 ⁽³⁾	0.38943 ⁽⁶⁾	0.35877 ⁽⁴⁾
			$\hat{\theta}$	0.17201 ⁽⁶⁾	0.20431 ⁽⁷⁾	0.13012 ⁽¹⁾	0.13860 ⁽²⁾	0.20509 ⁽⁸⁾	0.14233 ⁽⁴⁾	0.14966 ⁽⁵⁾	0.14228 ⁽³⁾
\hat{a}			0.18702 ⁽⁶⁾	0.22516 ⁽⁸⁾	0.11276 ⁽¹⁾	0.14184 ⁽²⁾	0.22124 ⁽⁷⁾	0.16230 ⁽⁵⁾	0.15764 ⁽³⁾	0.15905 ⁽⁴⁾	
MSE		$\hat{\beta}$	0.19686 ⁽⁵⁾	0.22092 ⁽⁶⁾	0.13575 ⁽¹⁾	0.16596 ⁽²⁾	0.22831 ⁽⁸⁾	0.18188 ⁽³⁾	0.22728 ⁽⁷⁾	0.19581 ⁽⁴⁾	
		$\hat{\theta}$	0.08375 ⁽⁶⁾	0.11051 ⁽⁷⁾	0.04891 ⁽¹⁾	0.05692 ⁽²⁾	0.11165 ⁽⁸⁾	0.05764 ⁽³⁾	0.06548 ⁽⁵⁾	0.05892 ⁽⁴⁾	
		\hat{a}	0.05577 ⁽⁶⁾	0.07600 ⁽⁸⁾	0.02255 ⁽¹⁾	0.03381 ⁽²⁾	0.07377 ⁽⁷⁾	0.04469 ⁽⁵⁾	0.04187 ⁽⁴⁾	0.04159 ⁽³⁾	
MRE		$\hat{\beta}$	0.17243 ⁽⁵⁾	0.18622 ⁽⁷⁾	0.14094 ⁽¹⁾	0.15612 ⁽²⁾	0.18922 ⁽⁸⁾	0.16399 ⁽³⁾	0.18544 ⁽⁶⁾	0.17084 ⁽⁴⁾	
		$\hat{\theta}$	1.72010 ⁽⁶⁾	2.04312 ⁽⁷⁾	1.30119 ⁽¹⁾	1.38603 ⁽²⁾	2.05089 ⁽⁸⁾	1.42328 ⁽⁴⁾	1.49663 ⁽⁵⁾	1.42281 ⁽³⁾	
		\hat{a}	0.37404 ⁽⁶⁾	0.45033 ⁽⁸⁾	0.22552 ⁽¹⁾	0.28368 ⁽²⁾	0.44248 ⁽⁷⁾	0.32459 ⁽⁵⁾	0.31527 ⁽³⁾	0.31809 ⁽⁴⁾	
$\sum Ranks$			51 ⁽⁶⁾	65 ⁽⁷⁾	9 ⁽¹⁾	18 ⁽²⁾	69 ⁽⁸⁾	35 ⁽⁴⁾	44 ⁽⁵⁾	33 ⁽³⁾	
80		BIAS	$\hat{\beta}$	0.28501 ⁽⁴⁾	0.33290 ⁽⁸⁾	0.23477 ⁽¹⁾	0.25361 ⁽²⁾	0.32441 ⁽⁷⁾	0.27238 ⁽³⁾	0.31394 ⁽⁶⁾	0.29088 ⁽⁵⁾
			$\hat{\theta}$	0.12397 ⁽⁶⁾	0.15818 ⁽⁸⁾	0.08891 ⁽¹⁾	0.09356 ⁽²⁾	0.15486 ⁽⁷⁾	0.10617 ⁽⁴⁾	0.11480 ⁽⁵⁾	0.09754 ⁽³⁾
	\hat{a}		0.14039 ⁽⁶⁾	0.18772 ⁽⁸⁾	0.08611 ⁽¹⁾	0.10338 ⁽²⁾	0.17818 ⁽⁷⁾	0.12315 ⁽⁵⁾	0.11890 ⁽⁴⁾	0.11521 ⁽³⁾	
	MSE	$\hat{\beta}$	0.12825 ⁽⁴⁾	0.16590 ⁽⁸⁾	0.08740 ⁽¹⁾	0.10191 ⁽²⁾	0.15756 ⁽⁷⁾	0.11618 ⁽³⁾	0.15238 ⁽⁶⁾	0.13291 ⁽⁵⁾	
		$\hat{\theta}$	0.04416 ⁽⁶⁾	0.06739 ⁽⁸⁾	0.02085 ⁽¹⁾	0.02425 ⁽²⁾	0.06467 ⁽⁷⁾	0.03041 ⁽⁴⁾	0.03789 ⁽⁵⁾	0.02543 ⁽³⁾	
		\hat{a}	0.03365 ⁽⁶⁾	0.05563 ⁽⁸⁾	0.01272 ⁽¹⁾	0.01765 ⁽²⁾	0.05123 ⁽⁷⁾	0.02705 ⁽⁵⁾	0.02487 ⁽⁴⁾	0.02211 ⁽³⁾	
	MRE	$\hat{\beta}$	0.13572 ⁽⁴⁾	0.15852 ⁽⁸⁾	0.11179 ⁽¹⁾	0.12077 ⁽²⁾	0.15448 ⁽⁷⁾	0.12970 ⁽³⁾	0.14949 ⁽⁶⁾	0.13851 ⁽⁵⁾	
		$\hat{\theta}$	1.23966 ⁽⁶⁾	1.58182 ⁽⁸⁾	0.88913 ⁽¹⁾	0.93555 ⁽²⁾	1.54855 ⁽⁷⁾	1.06169 ⁽⁴⁾	1.14797 ⁽⁵⁾	0.97543 ⁽³⁾	
		\hat{a}	0.28078 ⁽⁶⁾	0.37543 ⁽⁸⁾	0.17222 ⁽¹⁾	0.20676 ⁽²⁾	0.35637 ⁽⁷⁾	0.24629 ⁽⁵⁾	0.23780 ⁽⁴⁾	0.23043 ⁽³⁾	
	$\sum Ranks$			48 ⁽⁶⁾	72 ⁽⁸⁾	9 ⁽¹⁾	18 ⁽²⁾	63 ⁽⁷⁾	36 ⁽⁴⁾	45 ⁽⁵⁾	33 ⁽³⁾
	120	BIAS	$\hat{\beta}$	0.22809 ⁽⁴⁾	0.26757 ⁽⁷⁾	0.19568 ⁽¹⁾	0.20397 ⁽²⁾	0.27004 ⁽⁸⁾	0.22221 ⁽³⁾	0.25442 ⁽⁶⁾	0.24035 ⁽⁵⁾
			$\hat{\theta}$	0.09017 ⁽⁶⁾	0.12038 ⁽⁸⁾	0.06957 ⁽¹⁾	0.06978 ⁽²⁾	0.11980 ⁽⁷⁾	0.08069 ⁽⁴⁾	0.08668 ⁽⁵⁾	0.07263 ⁽³⁾
\hat{a}			0.10600 ⁽⁶⁾	0.14909 ⁽⁷⁾	0.07074 ⁽¹⁾	0.08047 ⁽²⁾	0.15022 ⁽⁸⁾	0.09961 ⁽⁵⁾	0.09573 ⁽⁴⁾	0.08727 ⁽³⁾	
MSE		$\hat{\beta}$	0.08267 ⁽⁴⁾	0.11131 ⁽⁷⁾	0.06029 ⁽¹⁾	0.06598 ⁽²⁾	0.11305 ⁽⁸⁾	0.07837 ⁽³⁾	0.10241 ⁽⁶⁾	0.09202 ⁽⁵⁾	
		$\hat{\theta}$	0.02147 ⁽⁶⁾	0.03973 ⁽⁸⁾	0.01214 ⁽²⁾	0.01157 ⁽¹⁾	0.03730 ⁽⁷⁾	0.01512 ⁽⁴⁾	0.01984 ⁽⁵⁾	0.01253 ⁽³⁾	
		\hat{a}	0.01956 ⁽⁶⁾	0.03754 ⁽⁷⁾	0.00861 ⁽¹⁾	0.01040 ⁽²⁾	0.03770 ⁽⁸⁾	0.01741 ⁽⁵⁾	0.01582 ⁽⁴⁾	0.01238 ⁽³⁾	
MRE		$\hat{\beta}$	0.10861 ⁽⁴⁾	0.12741 ⁽⁷⁾	0.09318 ⁽¹⁾	0.09713 ⁽²⁾	0.12859 ⁽⁸⁾	0.10581 ⁽³⁾	0.12115 ⁽⁶⁾	0.11445 ⁽⁵⁾	
		$\hat{\theta}$	0.90171 ⁽⁶⁾	1.20380 ⁽⁸⁾	0.69572 ⁽¹⁾	0.69782 ⁽²⁾	1.19803 ⁽⁷⁾	0.80686 ⁽⁴⁾	0.86682 ⁽⁵⁾	0.72632 ⁽³⁾	
		\hat{a}	0.21199 ⁽⁶⁾	0.29817 ⁽⁷⁾	0.14147 ⁽¹⁾	0.16093 ⁽²⁾	0.30044 ⁽⁸⁾	0.19923 ⁽⁵⁾	0.19147 ⁽⁴⁾	0.17454 ⁽³⁾	
$\sum Ranks$			48 ⁽⁶⁾	66 ⁽⁷⁾	10 ⁽¹⁾	17 ⁽²⁾	69 ⁽⁸⁾	36 ⁽⁴⁾	45 ⁽⁵⁾	33 ⁽³⁾	
200		BIAS	$\hat{\beta}$	0.17066 ⁽⁴⁾	0.20566 ⁽⁷⁾	0.14763 ⁽¹⁾	0.15401 ⁽²⁾	0.20861 ⁽⁸⁾	0.16913 ⁽³⁾	0.19246 ⁽⁶⁾	0.18538 ⁽⁵⁾
			$\hat{\theta}$	0.06290 ⁽⁶⁾	0.08330 ⁽⁷⁾	0.04858 ⁽¹⁾	0.04887 ⁽²⁾	0.08401 ⁽⁸⁾	0.05793 ⁽⁴⁾	0.06124 ⁽⁵⁾	0.05331 ⁽³⁾
	\hat{a}		0.07600 ⁽⁶⁾	0.11242 ⁽⁸⁾	0.05367 ⁽¹⁾	0.05936 ⁽²⁾	0.11078 ⁽⁷⁾	0.07142 ⁽⁵⁾	0.06959 ⁽⁴⁾	0.06428 ⁽³⁾	
	MSE	$\hat{\beta}$	0.04626 ⁽⁴⁾	0.06738 ⁽⁷⁾	0.03406 ⁽¹⁾	0.03755 ⁽²⁾	0.06906 ⁽⁸⁾	0.04562 ⁽³⁾	0.05946 ⁽⁶⁾	0.05432 ⁽⁵⁾	
		$\hat{\theta}$	0.00894 ⁽⁶⁾	0.01607 ⁽⁷⁾	0.00470 ⁽¹⁾	0.00475 ⁽²⁾	0.01658 ⁽⁸⁾	0.00694 ⁽⁴⁾	0.00861 ⁽⁵⁾	0.00536 ⁽³⁾	
		\hat{a}	0.00966 ⁽⁶⁾	0.02200 ⁽⁸⁾	0.00465 ⁽¹⁾	0.00553 ⁽²⁾	0.02127 ⁽⁷⁾	0.00861 ⁽⁵⁾	0.00807 ⁽⁴⁾	0.00645 ⁽³⁾	
	MRE	$\hat{\beta}$	0.08127 ⁽⁴⁾	0.09793 ⁽⁷⁾	0.07030 ⁽¹⁾	0.07334 ⁽²⁾	0.09934 ⁽⁸⁾	0.08054 ⁽³⁾	0.09165 ⁽⁶⁾	0.08828 ⁽⁵⁾	
		$\hat{\theta}$	0.62897 ⁽⁶⁾	0.83304 ⁽⁷⁾	0.48578 ⁽¹⁾	0.48866 ⁽²⁾	0.84010 ⁽⁸⁾	0.57932 ⁽⁴⁾	0.61238 ⁽⁵⁾	0.53313 ⁽³⁾	
		\hat{a}	0.15199 ⁽⁶⁾	0.22485 ⁽⁸⁾	0.10734 ⁽¹⁾	0.11872 ⁽²⁾	0.22156 ⁽⁷⁾	0.14285 ⁽⁵⁾	0.13917 ⁽⁴⁾	0.12856 ⁽³⁾	
	$\sum Ranks$			48 ⁽⁶⁾	66 ⁽⁷⁾	9 ⁽¹⁾	18 ⁽²⁾	69 ⁽⁸⁾	36 ⁽⁴⁾	45 ⁽⁵⁾	33 ⁽³⁾

The observations from the MOWE model are simulated from Eq. (8), where U is a uniform random variable in the interval $(0, 1)$. We generate $N = 5,000$ random samples x_1, \dots, x_n (for $n = 30, 50, 80, 120$, and 200) from the MOWE model with $\beta = \{0.2, 1.0, 2.1\}$, $\theta = \{0.1, 1.1, 1.2\}$ and $a = \{0.5, 1.5, 4.7\}$. We use R codes (R Core Team, 2020, version 4.0.3) for the simulations

and the `nlminb` function in the `stats` package (R Core Team, 2020, version 4.0.3).

We estimate its parameters for some parameter combinations and sample sizes, and calculate the MSEs and MREs of the estimates. Four out of twenty-seven simulated outcomes are reported in Tables 1–4, whose numbers in each row have superscripts giving

Table 4
Simulation results for $\beta = 0.2, \theta = 0.1, \alpha = 1.5$.

<i>n</i>	Est.	Est. Par.	WLSE	OLSE	MLE	MPSE	CVME	ADE	RADE	PCE	
30	BIAS	$\hat{\beta}$	0.03025 ⁽⁴⁾	0.03200 ⁽⁵⁾	0.01860 ⁽³⁾	0.01381 ⁽¹⁾	0.03468 ⁽⁶⁾	0.01755 ⁽²⁾	0.03598 ⁽⁷⁾	0.13825 ⁽⁸⁾	
		$\hat{\theta}$	0.04936 ⁽⁴⁾	0.05359 ⁽⁶⁾	0.02367 ⁽¹⁾	0.02414 ⁽²⁾	0.05040 ⁽⁵⁾	0.02584 ⁽³⁾	0.05373 ⁽⁷⁾	0.08650 ⁽⁸⁾	
		$\hat{\alpha}$	0.53838 ⁽⁶⁾	0.42161 ⁽⁵⁾	0.14593 ⁽¹⁾	0.17715 ⁽²⁾	0.38506 ⁽⁴⁾	0.28715 ⁽³⁾	0.75793 ⁽⁷⁾	0.90701 ⁽⁸⁾	
	MSE	$\hat{\beta}$	0.00157 ⁽⁴⁾	0.00172 ⁽⁵⁾	0.00109 ⁽³⁾	0.00059 ⁽¹⁾	0.00212 ⁽⁶⁾	0.00097 ⁽²⁾	0.00226 ⁽⁷⁾	0.03846 ⁽⁸⁾	
		$\hat{\theta}$	0.00418 ⁽⁴⁾	0.00527 ⁽⁷⁾	0.00175 ⁽¹⁾	0.00191 ⁽²⁾	0.00432 ⁽⁵⁾	0.00209 ⁽³⁾	0.00479 ⁽⁶⁾	0.01219 ⁽⁸⁾	
		$\hat{\alpha}$	0.54954 ⁽⁶⁾	0.45582 ⁽⁵⁾	0.10795 ⁽¹⁾	0.17236 ⁽²⁾	0.40190 ⁽⁴⁾	0.28395 ⁽³⁾	0.76507 ⁽⁷⁾	1.00780 ⁽⁸⁾	
	MRE	$\hat{\beta}$	0.15123 ⁽⁴⁾	0.16001 ⁽⁵⁾	0.09298 ⁽³⁾	0.06904 ⁽¹⁾	0.17339 ⁽⁶⁾	0.08777 ⁽²⁾	0.17990 ⁽⁷⁾	0.69123 ⁽⁸⁾	
		$\hat{\theta}$	0.49356 ⁽⁴⁾	0.53593 ⁽⁶⁾	0.23671 ⁽¹⁾	0.24136 ⁽²⁾	0.50400 ⁽⁵⁾	0.25842 ⁽³⁾	0.53728 ⁽⁷⁾	0.86499 ⁽⁸⁾	
		$\hat{\alpha}$	0.35892 ⁽⁶⁾	0.28107 ⁽⁵⁾	0.09729 ⁽¹⁾	0.11810 ⁽²⁾	0.25671 ⁽⁴⁾	0.19143 ⁽³⁾	0.50529 ⁽⁷⁾	0.60468 ⁽⁸⁾	
	$\sum Ranks$			42 ⁽⁴⁾	49 ⁽⁶⁾	15 ^(1.5)	15 ^(1.5)	45 ⁽⁵⁾	24 ⁽³⁾	62 ⁽⁷⁾	72 ⁽⁸⁾
	50	BIAS	$\hat{\beta}$	0.02346 ⁽⁴⁾	0.02455 ⁽⁵⁾	0.01016 ⁽³⁾	0.00744 ⁽¹⁾	0.02598 ⁽⁶⁾	0.00912 ⁽²⁾	0.02720 ⁽⁷⁾	0.12833 ⁽⁸⁾
			$\hat{\theta}$	0.03857 ⁽⁴⁾	0.03994 ⁽⁶⁾	0.01233 ⁽²⁾	0.01221 ⁽¹⁾	0.03980 ⁽⁵⁾	0.01330 ⁽³⁾	0.04218 ⁽⁷⁾	0.07044 ⁽⁸⁾
$\hat{\alpha}$			0.44584 ⁽⁶⁾	0.35853 ⁽⁵⁾	0.05855 ⁽¹⁾	0.08010 ⁽²⁾	0.34030 ⁽⁴⁾	0.16441 ⁽³⁾	0.67734 ⁽⁷⁾	0.96533 ⁽⁸⁾	
MSE		$\hat{\beta}$	0.00089 ⁽⁴⁾	0.00098 ⁽⁵⁾	0.00046 ⁽³⁾	0.00026 ⁽¹⁾	0.00113 ⁽⁶⁾	0.00039 ⁽²⁾	0.00122 ⁽⁷⁾	0.03104 ⁽⁸⁾	
		$\hat{\theta}$	0.00246 ⁽⁴⁾	0.00271 ⁽⁶⁾	0.00065 ⁽¹⁾	0.00067 ⁽²⁾	0.00266 ⁽⁵⁾	0.00079 ⁽³⁾	0.00294 ⁽⁷⁾	0.00751 ⁽⁸⁾	
		$\hat{\alpha}$	0.42044 ⁽⁶⁾	0.36665 ⁽⁵⁾	0.03891 ⁽¹⁾	0.06592 ⁽²⁾	0.33668 ⁽⁴⁾	0.14698 ⁽³⁾	0.61742 ⁽⁷⁾	1.05429 ⁽⁸⁾	
MRE		$\hat{\beta}$	0.11730 ⁽⁴⁾	0.12276 ⁽⁵⁾	0.05079 ⁽³⁾	0.03720 ⁽¹⁾	0.12988 ⁽⁶⁾	0.04562 ⁽²⁾	0.13599 ⁽⁷⁾	0.64165 ⁽⁸⁾	
		$\hat{\theta}$	0.38569 ⁽⁴⁾	0.39941 ⁽⁶⁾	0.12332 ⁽²⁾	0.12206 ⁽¹⁾	0.39804 ⁽⁵⁾	0.13303 ⁽³⁾	0.42178 ⁽⁷⁾	0.70444 ⁽⁸⁾	
		$\hat{\alpha}$	0.29723 ⁽⁶⁾	0.23902 ⁽⁵⁾	0.03904 ⁽¹⁾	0.05340 ⁽²⁾	0.22687 ⁽⁴⁾	0.10961 ⁽³⁾	0.45156 ⁽⁷⁾	0.64355 ⁽⁸⁾	
$\sum Ranks$			42 ⁽⁴⁾	48 ⁽⁶⁾	17 ⁽²⁾	13 ⁽¹⁾	45 ⁽⁵⁾	24 ⁽³⁾	63 ⁽⁷⁾	72 ⁽⁸⁾	
80		BIAS	$\hat{\beta}$	0.01801 ⁽⁴⁾	0.01975 ⁽⁵⁾	0.00427 ⁽³⁾	0.00328 ⁽¹⁾	0.02016 ⁽⁶⁾	0.00394 ⁽²⁾	0.02116 ⁽⁷⁾	0.12687 ⁽⁸⁾
			$\hat{\theta}$	0.03052 ⁽⁴⁾	0.03325 ⁽⁷⁾	0.00512 ⁽²⁾	0.00494 ⁽¹⁾	0.03163 ⁽⁵⁾	0.00565 ⁽³⁾	0.03307 ⁽⁶⁾	0.05692 ⁽⁸⁾
	$\hat{\alpha}$		0.38097 ⁽⁶⁾	0.31683 ⁽⁵⁾	0.01836 ⁽¹⁾	0.01966 ⁽²⁾	0.31273 ⁽⁴⁾	0.07742 ⁽³⁾	0.59552 ⁽⁷⁾	0.90925 ⁽⁸⁾	
	MSE	$\hat{\beta}$	0.00054 ⁽⁴⁾	0.00063 ⁽⁵⁾	0.00015 ⁽³⁾	0.00010 ⁽¹⁾	0.00067 ⁽⁶⁾	0.00013 ⁽²⁾	0.00074 ⁽⁷⁾	0.03151 ⁽⁸⁾	
		$\hat{\theta}$	0.00149 ⁽⁴⁾	0.00191 ⁽⁷⁾	0.00021 ^(1.5)	0.00021 ^(1.5)	0.00163 ⁽⁵⁾	0.00026 ⁽³⁾	0.00179 ⁽⁶⁾	0.00496 ⁽⁸⁾	
		$\hat{\alpha}$	0.32667 ⁽⁶⁾	0.30282 ⁽⁵⁾	0.00978 ⁽¹⁾	0.01168 ⁽²⁾	0.29221 ⁽⁴⁾	0.06292 ⁽³⁾	0.49183 ⁽⁷⁾	0.95310 ⁽⁸⁾	
	MRE	$\hat{\beta}$	0.09006 ⁽⁴⁾	0.09877 ⁽⁵⁾	0.02136 ⁽³⁾	0.01640 ⁽¹⁾	0.10082 ⁽⁶⁾	0.01969 ⁽²⁾	0.10578 ⁽⁷⁾	0.63435 ⁽⁸⁾	
		$\hat{\theta}$	0.30518 ⁽⁴⁾	0.33249 ⁽⁷⁾	0.05123 ⁽²⁾	0.04939 ⁽¹⁾	0.31626 ⁽⁵⁾	0.05649 ⁽³⁾	0.33073 ⁽⁶⁾	0.56920 ⁽⁸⁾	
		$\hat{\alpha}$	0.25398 ⁽⁶⁾	0.21122 ⁽⁵⁾	0.01224 ⁽¹⁾	0.01311 ⁽²⁾	0.20849 ⁽⁴⁾	0.05161 ⁽³⁾	0.39701 ⁽⁷⁾	0.60617 ⁽⁸⁾	
	$\sum Ranks$			42 ⁽⁴⁾	51 ⁽⁶⁾	17.5 ⁽²⁾	12.5 ⁽¹⁾	45 ⁽⁵⁾	24 ⁽³⁾	60 ⁽⁷⁾	72 ⁽⁸⁾
	120	BIAS	$\hat{\beta}$	0.01499 ⁽⁴⁾	0.01631 ⁽⁶⁾	0.00166 ⁽³⁾	0.00121 ⁽¹⁾	0.01577 ⁽⁵⁾	0.00141 ⁽²⁾	0.01680 ⁽⁷⁾	0.12266 ⁽⁸⁾
			$\hat{\theta}$	0.02522 ⁽⁴⁾	0.02746 ⁽⁷⁾	0.00195 ⁽²⁾	0.00165 ⁽¹⁾	0.02577 ⁽⁵⁾	0.00200 ⁽³⁾	0.02667 ⁽⁶⁾	0.04851 ⁽⁸⁾
$\hat{\alpha}$			0.33734 ⁽⁶⁾	0.30433 ⁽⁵⁾	0.00537 ⁽²⁾	0.00524 ⁽¹⁾	0.29507 ⁽⁴⁾	0.02703 ⁽³⁾	0.49719 ⁽⁷⁾	0.83923 ⁽⁸⁾	
MSE		$\hat{\beta}$	0.00036 ⁽⁴⁾	0.00042 ⁽⁶⁾	0.00005 ⁽³⁾	0.00003 ⁽¹⁾	0.00040 ⁽⁵⁾	0.00004 ⁽²⁾	0.00046 ⁽⁷⁾	0.03193 ⁽⁸⁾	
		$\hat{\theta}$	0.00103 ⁽⁴⁾	0.00120 ⁽⁷⁾	0.00006 ⁽²⁾	0.00005 ⁽¹⁾	0.00108 ⁽⁵⁾	0.00008 ⁽³⁾	0.00114 ⁽⁶⁾	0.00363 ⁽⁸⁾	
		$\hat{\alpha}$	0.26320 ⁽⁴⁾	0.27959 ⁽⁶⁾	0.00247 ⁽²⁾	0.00224 ⁽¹⁾	0.26513 ⁽⁵⁾	0.02024 ⁽³⁾	0.36163 ⁽⁷⁾	0.83621 ⁽⁸⁾	
MRE		$\hat{\beta}$	0.07494 ⁽⁴⁾	0.08155 ⁽⁶⁾	0.00830 ⁽³⁾	0.00604 ⁽¹⁾	0.07886 ⁽⁵⁾	0.00706 ⁽²⁾	0.08402 ⁽⁷⁾	0.61332 ⁽⁸⁾	
		$\hat{\theta}$	0.25219 ⁽⁴⁾	0.27462 ⁽⁷⁾	0.01953 ⁽²⁾	0.01649 ⁽¹⁾	0.25774 ⁽⁵⁾	0.01997 ⁽³⁾	0.26674 ⁽⁶⁾	0.48511 ⁽⁸⁾	
		$\hat{\alpha}$	0.22490 ⁽⁶⁾	0.20288 ⁽⁵⁾	0.00358 ⁽²⁾	0.00349 ⁽¹⁾	0.19671 ⁽⁴⁾	0.01802 ⁽³⁾	0.33146 ⁽⁷⁾	0.55948 ⁽⁸⁾	
$\sum Ranks$			40 ⁽⁴⁾	55 ⁽⁶⁾	21 ⁽²⁾	9 ⁽¹⁾	43 ⁽⁵⁾	24 ⁽³⁾	60 ⁽⁷⁾	72 ⁽⁸⁾	
200		BIAS	$\hat{\beta}$	0.01142 ⁽⁴⁾	0.01248 ⁽⁶⁾	0.00030 ⁽³⁾	0.00023 ⁽¹⁾	0.01243 ⁽⁵⁾	0.00025 ⁽²⁾	0.01323 ⁽⁷⁾	0.10550 ⁽⁸⁾
			$\hat{\theta}$	0.01936 ⁽⁴⁾	0.02124 ⁽⁶⁾	0.00030 ⁽¹⁾	0.00032 ⁽²⁾	0.02090 ⁽⁵⁾	0.00035 ⁽³⁾	0.02131 ⁽⁷⁾	0.03986 ⁽⁸⁾
	$\hat{\alpha}$		0.29508 ⁽⁶⁾	0.27993 ⁽⁴⁾	0.00035 ⁽¹⁾	0.00064 ⁽²⁾	0.28607 ⁽⁵⁾	0.00562 ⁽³⁾	0.40847 ⁽⁷⁾	0.75140 ⁽⁸⁾	
	MSE	$\hat{\beta}$	0.00021 ⁽⁴⁾	0.00025 ^(5.5)	0.00001 ^(2.5)	0.00000 ⁽¹⁾	0.00025 ^(5.5)	0.00001 ^(2.5)	0.00028 ⁽⁷⁾	0.02512 ⁽⁸⁾	
		$\hat{\theta}$	0.00060 ⁽⁴⁾	0.00075 ⁽⁷⁾	0.00001 ⁽²⁾	0.00001 ⁽²⁾	0.00072 ^(5.5)	0.00001 ⁽²⁾	0.00072 ^(5.5)	0.00241 ⁽⁸⁾	
		$\hat{\alpha}$	0.20688 ⁽⁴⁾	0.24122 ⁽⁵⁾	0.00004 ⁽¹⁾	0.00006 ⁽²⁾	0.24673 ⁽⁶⁾	0.00351 ⁽³⁾	0.25467 ⁽⁷⁾	0.68867 ⁽⁸⁾	
	MRE	$\hat{\beta}$	0.05709 ⁽⁴⁾	0.06241 ⁽⁶⁾	0.00152 ⁽³⁾	0.00113 ⁽¹⁾	0.06216 ⁽⁵⁾	0.00123 ⁽²⁾	0.06613 ⁽⁷⁾	0.52750 ⁽⁸⁾	
		$\hat{\theta}$	0.19363 ⁽⁴⁾	0.21242 ⁽⁶⁾	0.00303 ⁽¹⁾	0.00316 ⁽²⁾	0.20899 ⁽⁵⁾	0.00347 ⁽³⁾	0.21308 ⁽⁷⁾	0.39861 ⁽⁸⁾	
		$\hat{\alpha}$	0.19672 ⁽⁶⁾	0.18662 ⁽⁴⁾	0.00024 ⁽¹⁾	0.00043 ⁽²⁾	0.19071 ⁽⁵⁾	0.00375 ⁽³⁾	0.27231 ⁽⁷⁾	0.50094 ⁽⁸⁾	
	$\sum Ranks$			40 ⁽⁴⁾	49.5 ⁽⁶⁾	15.5 ⁽²⁾	15 ⁽¹⁾	47 ⁽⁵⁾	23.5 ⁽³⁾	61.5 ⁽⁷⁾	72 ⁽⁸⁾

the ranks of the estimates among all methods, and $\sum Ranks$ denotes the partial sum of the ranks.

Table 5 provides the partial and overall ranks of the estimates, thus indicating that the MLEs outperform all other estimates for the MOWE distribution with an overall score of 230.

6. Modeling biological data

The applicability of a sub-model of the new family is proved empirically in modeling two COVID-19 data sets.

The first set refers to 36 COVID-19 mortality rates in Canada: 1.5157, 1.5806, 1.9048, 2.1901, 2.4141, 2.4946, 2.5261, 2.6029, 2.7704, 2.7957, 2.8349, 2.8636, 2.9078, 3.0914, 3.1091, 3.1091, 3.1444, 3.1348, 3.2110, 3.2135, 3.2218, 3.2823, 3.3592, 3.3769, 3.3825, 3.5146, 3.6346, 3.6426, 3.8594, 4.0480, 4.1685, 4.2202, 4.2781, 4.9274, 4.9378, 6.8686. The second set refers to 53 COVID-19 survival times of patients in critical conditions in China in the first two months of 2020. The times measured from the admission to the hospital until death are: 0.054, 0.064, 0.087, 0.087, 0.235, 0.352, 0.364, 0.421, 0.437, 0.458, 0.479, 0.548,

Table 5
Partial and overall ranks of all estimates for some combinations of θ .

θ^T	n	WLSE	OLSE	MLE	MPSE	CVME	ADE	RADE	PCE
$(\beta = 0.2, \theta = 0.1, a = 0.5)$	30	3	5	4	1	6	2	7	8
	50	4	5	3	1	6	2	7	8
	80	4	5	3	1	6	2	7	8
	120	4	5	3	1	6.5	2	6.5	8
	200	4	7	3	1	6	2	5	8
$(\beta = 0.2, \theta = 0.1, a = 1.5)$	30	4	6	1.5	1.5	5	3	7	8
	50	4	6	2	1	5	3	7	8
	80	4	6	2	1	5	3	7	8
	120	4	6	2	1	5	3	7	8
	200	4	6	2	1	5	3	7	8
$(\beta = 0.2, \theta = 0.1, a = 4.7)$	30	4	6	1.5	1.5	5	3	7	8
	50	4	6	2	1	5	3	7	8
	80	4	5	1	2	6	3	7	8
	120	4	6	2	1	5	3	7	8
	200	4	6	3	1	5	2	7	8
$(\beta = 0.2, \theta = 1.1, a = 0.5)$	30	3	6	4	1	7	2	5	8
	50	3	6	5	1	7	2	4	8
	80	3	5	7	1	6	2	4	8
	120	3	4.5	7	1	6	2	4.5	8
	200	3	6	7	1	5	2	4	8
$(\beta = 0.2, \theta = 1.1, a = 1.5)$	30	4	6	2	1	7	3	5	8
	50	4	7	1	2	6	3	5	8
	80	4	6	2	1	7	3	5	8
	120	4	6	2	1	7	3	5	8
	200	4	7	1	2	6	3	5	8
$(\beta = 0.2, \theta = 1.1, a = 4.7)$	30	4	5	2	1	6.5	3	6.5	8
	50	4	5	2	1	7	3	6	8
	80	4	7	1	2	6	3	5	8
	120	4	6	2	1	7	3	5	8
	200	4	7	2	1	6	3	5	8
$(\beta = 0.2, \theta = 1.2, a = 0.5)$	30	3	6	4	1	7	2	5	8
	50	3	7	5.5	1	5.5	2	4	8
	80	3	5	7	1	6	2	4	8
	120	3	5	7	1	6	2	4	8
	200	3	6	7	1	5	2	4	8
$(\beta = 0.2, \theta = 1.2, a = 1.5)$	30	4	5	2	1	7	3	6	8
	50	4	6	2	1	7	3	5	8
	80	4	6	2	1	7	3	5	8
	120	4	5.5	2	1	7	3	5.5	8
	200	4	7	2	1	6	3	5	8
$(\beta = 0.2, \theta = 1.2, a = 4.7)$	30	4	5	2	1	7	3	6	8
	50	4	7	2	1	6	3	5	8
	80	4	5	1	2	7	3	6	8
	120	4	6	2	1	7	3	5	8
	200	4	6	2	1	7	3	5	8
$(\beta = 1.0, \theta = 0.1, a = 0.5)$	30	4.5	6	1	2	7	3	4.5	8
	50	4.5	7	1	2	6	3	4.5	8
	80	4.5	7	1	2	6	3	4.5	8
	120	5	6	1	2	7	3	4	8
	200	5	7	1	2	6	3	4	8
$(\beta = 1.0, \theta = 0.1, a = 1.5)$	30	5	7	1	2	6	3	4	8
	50	4.5	7	1	2	6	3	4.5	8
	80	4	7	1	2	6	3	5	8
	120	5	7	1	2	6	3	4	8
	200	5	6	1	2	7	3	4	8
$(\beta = 1.0, \theta = 0.1, a = 4.7)$	30	5	7	1	2	6	3	4	8
	50	5	7	1	2	6	3	4	8
	80	4.5	7	1	2	6	3	4.5	8
	120	4	7	1	2	6	3	5	8
	200	4	7.5	1	2	6	3	5	7.5
$(\beta = 1.0, \theta = 1.1, a = 0.5)$	30	4	8	1	2	5	3	6	7
	50	4	7	1	2	5	3	6	8
	80	4	6	1	2	5	3	7	8
	120	4	5	1	2	6	3	7	8
	200	4	5	1	2	6	3	8	7
$(\beta = 1.0, \theta = 1.1, a = 1.5)$	30	4	7.5	1	2	5	3	6	7.5
	50	4	5	1	2	7	3	6	8
	80	4	6	1	2	5	3	7	8

(continued on next page)

Table 5 (continued)

θ^T	n	WLSE	OLSE	MLE	MPSE	CVME	ADE	RADE	PCE
	120	4	7	1	2	5	3	6	8
	200	4	6	1	2	5	3	7	8
$(\beta = 1.0, \theta = 1.1, a = 4.7)$	30	4	6	1	2	5	3	7	8
	50	4	6	1	2	5	3	7	8
	80	4	5	1	2	6	3	7	8
	120	4	6	2	1	5	3	7	8
	200	4	6	2	1	5	3	7	8
	30	4	6	1	2	5	3	7	8
$(\beta = 1.0, \theta = 1.2, a = 0.5)$	50	4	6	1	2	5	3	8	7
	80	4	6	1	2	5	3	7	8
	120	4	6	1	2	5	3	7	8
	200	4	7	1	2	5	3	8	6
	30	4	8	1	2	5	3	6	7
$(\beta = 1.0, \theta = 1.2, a = 1.5)$	50	4	6.5	1	2	5	3	8	6.5
	80	4	5	1	2	6	3	8	7
	120	4	6	1	2	5	3	8	7
	200	4	6	1	2	5	3	8	7
	30	4	5	2	1	6	3	8	7
$(\beta = 1.0, \theta = 1.2, a = 4.7)$	50	4	6	2	1	5	3	8	7
	80	4	5	2	1	6	3	7	8
	120	4	6	2	1	5	3	8	7
	200	4	6	2	1	5	3	7	8
	30	6	8	1	2	7	3	5	4
$(\beta = 2.1, \theta = 0.1, a = 0.5)$	50	6	7	1	2	8	4	5	3
	80	6	8	1	2	7	4	5	3
	120	6	7	1	2	8	4	5	3
	200	6	7	1	2	8	4	5	3
	30	6	7	1	2.5	8	2.5	5	4
$(\beta = 2.1, \theta = 0.1, a = 1.5)$	50	5	7	1	2	8	4	6	3
	80	6	7	1	2	8	4	5	3
	120	6	8	1	2	7	4	5	3
	200	5	8	1	2	7	4	6	3
	30	6	8	1	3	7	2	4.5	4.5
$(\beta = 2.1, \theta = 0.1, a = 4.7)$	50	6	7.5	1	2	7.5	3	4	5
	80	6	8	1	2	7	3	5	4
	120	6	8	1	2	7	4	5	3
	200	5	8	1	2	7	4	6	3
	30	6	8	1	5	4	2	7	3
$(\beta = 2.1, \theta = 1.1, a = 0.5)$	50	4	8	1	5	6	2	7	3
	80	5	7	1	4	6	2	8	3
	120	5	7	1	2	6	3	8	4
	200	5	7	1	2	6	3	8	4
	30	7	8	1	5	4	2	6	3
$(\beta = 2.1, \theta = 1.1, a = 1.5)$	50	4	8	1	5	6	2	7	3
	80	5	7	1	3	6	4	8	2
	120	5	7	1	4	6	3	8	2
	200	5	7	1	2	6	4	8	3
	30	6	8	1	4	5	2	7	3
$(\beta = 2.1, \theta = 1.1, a = 4.7)$	50	5	7	1	3	6	2	8	4
	80	5	7	1	3	6	2	8	4
	120	5	7	1	2	6	4	8	3
	200	5	7	1	2	6	3	8	4
	30	5	8	1	6	4	2	7	3
$(\beta = 2.1, \theta = 1.2, a = 0.5)$	50	5	8	1	3	6	2	7	4
	80	4	7	1	2	6	3	8	5
	120	5	7	1	2	6	4	8	3
	200	5	7	1	2	6	4	8	3
	30	5	8	1	4	6	2	7	3
$(\beta = 2.1, \theta = 1.2, a = 1.5)$	50	6	8	1	4	5	2	7	3
	80	5	7	1	4	6	2	8	3
	120	5	7	1	2	6	3	8	4
	200	5	7	1	2	6	4	8	3
	30	6	8	1.5	3	5	1.5	7	4
$(\beta = 2.1, \theta = 1.2, a = 4.7)$	50	6	7	1	4	5	2	8	3
	80	5	7	1	2.5	6	2.5	8	4
	120	5	7	1	2	6	3.5	8	3.5
	200	5	7	2	1	6	4	8	3
	\sum Ranks		599.5	881	230	260	808	389	836
Overall Rank		4	8	1	2	5	3	6	7

Table 6
Findings from the fitted distributions to the COVID-19 mortality rates.

Model	Par.	Estimates	(SEs)	AIC	CAIC	BIC	HQIC	W*	A*	K-S	K-S p-value
MOWE	$\hat{\beta}$	5.50558	(0.67807)	100.042	100.792	104.793	101.700	0.06008	0.34880	0.09773	0.88177
	$\hat{\theta}$	0.00279	(0.00094)								
	\hat{a}	0.09308	(0.01173)								
MOAPE	$\hat{\alpha}$	633804.1	(8388.73)	101.659	102.409	106.410	103.317	0.08217	0.45835	0.10992	0.77716
	$\hat{\lambda}$	1.88309	(0.32586)								
	$\hat{\theta}$	30.70694	(33.68783)								
TLOLLE	$\hat{\lambda}$	0.20969	(0.03824)	101.103	101.853	105.854	102.761	0.07514	0.43090	0.10240	0.84465
	\hat{a}	2.84413	(0.96418)								
	\hat{b}	2.02808	(1.55940)								
MONH	$\hat{\alpha}$	0.78838	(0.22230)	101.433	102.183	106.184	103.091	0.07280	0.40724	0.10341	0.83613
	$\hat{\lambda}$	4.30678	(4.48412)								
	$\hat{\theta}$	1539.80599	(2478.19647)								
BE	\hat{a}	15.14949	(10.67618)	101.991	102.741	106.742	103.649	0.09263	0.53929	0.10492	0.82298
	\hat{b}	2.21092	(1.95802)								
	$\hat{\lambda}$	0.78561	(0.38565)								
TGE	$\hat{\alpha}$	26.86514	(14.61735)	101.799	102.549	106.549	103.457	0.09225	0.54625	0.11143	0.76270
	$\hat{\lambda}$	1.31236	(0.17246)								
	$\hat{\theta}$	-0.65410	(0.34168)								
MOGE	$\hat{\alpha}$	31.58462	(24.81951)	101.064	101.814	105.815	102.722	0.08096	0.45622	0.10652	0.80864
	$\hat{\lambda}$	1.78986	(0.36869)								
	$\hat{\theta}$	9.09406	(13.42972)								
ME	$\hat{\alpha}$	1.83639	(1.47673)	103.578	104.869	109.912	105.789	0.08547	0.49574	0.10581	0.81504
	$\hat{\beta}$	5.41691	(9.80746)								
	$\hat{\gamma}$	17.07664	(17.24968)								
	$\hat{\lambda}$	1.04794	(0.56020)								
E	$\hat{\lambda}$	0.30473	(0.05078)	159.560	159.677	161.143	160.112	0.09950	0.57412	0.40970	0.00001

Table 7
Findings from the fitted distributions to the COVID-19 survival times.

Model	Par.	Estimates	(SEs)	AIC	CAIC	BIC	HQIC	W*	A*	K-S	K-S p-value
MOWE	$\hat{\beta}$	0.81250	(0.15573)	270.385	270.875	276.296	272.658	0.06537	0.41252	0.11296	0.50828
	$\hat{\theta}$	0.33504	(0.22823)								
	\hat{a}	0.07572	(0.02532)								
MOAPE	$\hat{\alpha}$	0.99990	(1.84449)	273.312	273.802	279.223	275.585	0.08339	0.49486	0.13065	0.32606
	$\hat{\lambda}$	0.12275	(0.04762)								
	$\hat{\theta}$	0.35277	(0.37871)								
TLOLLE	$\hat{\lambda}$	0.22384	(0.11090)	270.538	271.027	276.448	272.811	0.06146	0.39944	0.11682	0.46466
	\hat{a}	0.47642	(0.15364)								
	\hat{b}	2.21427	(1.02144)								
MONH	$\hat{\alpha}$	0.40812	(0.21923)	273.830	274.319	279.741	276.103	0.08558	0.51428	0.12036	0.42635
	$\hat{\lambda}$	2.54060	(7.71478)								
	$\hat{\theta}$	2.47644	(5.15816)								
BE	\hat{a}	0.69168	(0.12646)	272.840	273.330	278.751	275.113	0.07697	0.50460	0.13050	0.32739
	\hat{b}	1.00261	(2.99548)								
	$\hat{\lambda}$	0.16232	(0.51865)								
TGE	$\hat{\alpha}$	0.73423	(0.13934)	272.628	273.118	278.539	274.901	0.075175	0.48307	0.12597	0.36956
	$\hat{\lambda}$	0.15049	(0.04462)								
	$\hat{\theta}$	0.21354	(0.46487)								
MOGE	$\hat{\alpha}$	0.80648	(0.19856)	272.353	272.842	278.263	274.626	0.07403	0.46144	0.11861	0.44508
	$\hat{\lambda}$	0.13806	(0.04808)								
	$\hat{\theta}$	0.60440	(0.44413)								
ME	$\hat{\alpha}$	3.84904467	(6.14941)	274.365	275.198	282.246	277.396	0.07294	0.45987	0.12148	0.41461
	$\hat{\beta}$	823.54185	(383.2700)								
	$\hat{\gamma}$	0.43063	(0.12777)								
	$\hat{\lambda}$	0.03553	(0.08355)								
E	$\hat{\lambda}$	0.20892	(0.02869)	273.977	274.055	275.947	274.734	0.07751	0.50704	0.21143	0.01751

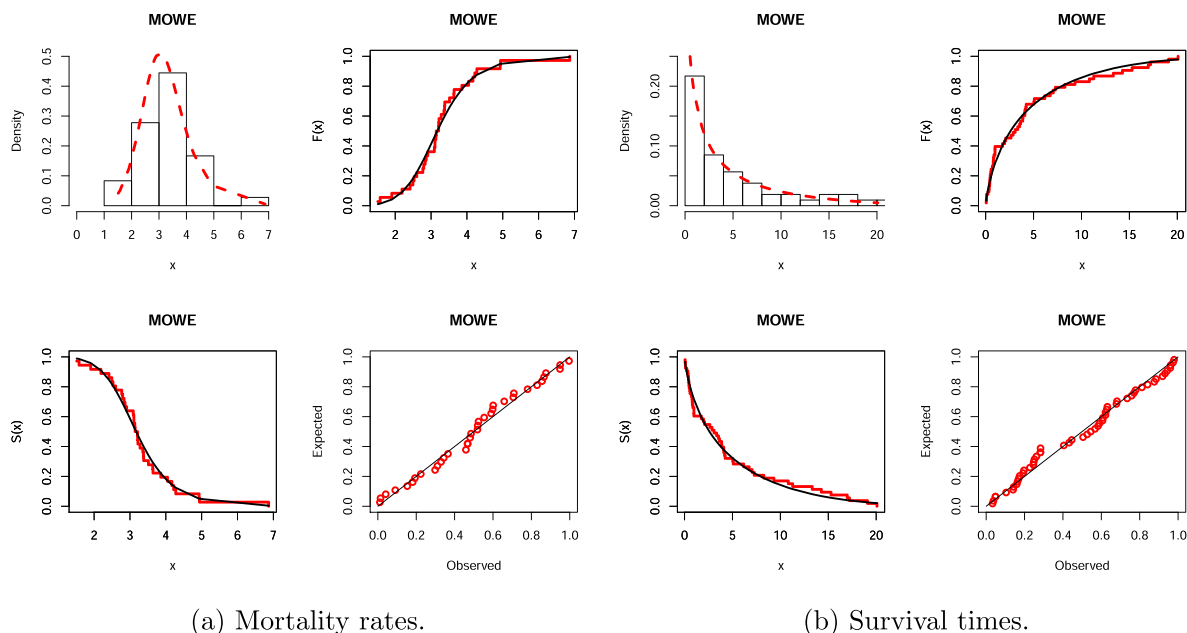


Fig. 2. Fitted functions for the MOWE model for the two data sets.

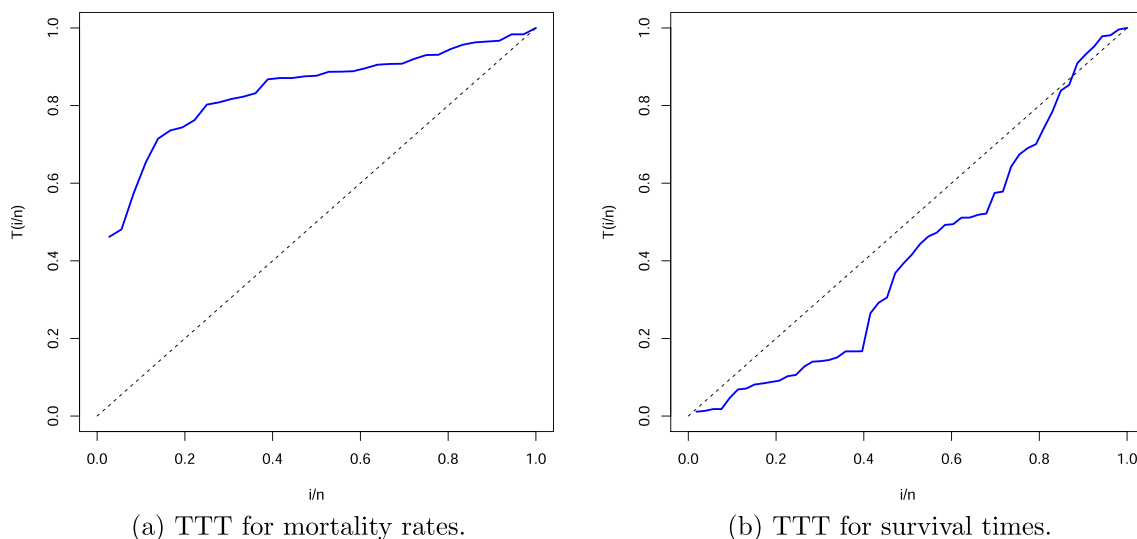


Fig. 3. TTT plots for the two analyzed data sets.

0.568, 0.704, 0.787, 0.796, 0.816, 0.865, 0.976, 0.976, 0.978, 1.756, 1.978, 2.089, 2.643, 2.869, 3.079, 3.348, 3.543, 3.646, 3.867, 3.890, 4.092, 4.093, 4.190, 4.237, 5.028, 5.083, 6.174, 6.743, 7.058, 7.274, 8.273, 9.324, 10.827, 11.282, 13.324, 14.278, 15.287, 16.978, 17.209, 19.092, 20.083. The two data sets were analyzed by Liu et al. (2021).

We compare the fits of the MOWE model and some other extensions of the exponential (E): the beta-E (BE) (Jones, 2004), Marshall–Olkin-generalized E (MOGE) (Ristic and Kundu, 2015), Marshall–Olkin-Nadarajah–Haghighi (MONH) (Lemonte et al., 2016), transmuted generalized-E (TGE) (Khan et al., 2017), modified-E (ME) (Rasekhi et al., 2017), Marshall–Olkin-alpha power E (MOAPE) (Nassar et al., 2019) and Topp–Leone-odd log-logistic E (TLOLE) (Afify et al., 2021a) distributions.

We adopt the information criterion (IC) measures: Akaike-IC (AIC), consistent Akaike-IC (CAIC), Hannan–Quinn IC (HQIC), Bayesian-IC (BIC), Cramér–Von Mises (W^*), Anderson–Darling (A^*), and Kolmogorov–Smirnov (K–S) (and K–S p -value).

The MLEs of the parameters from the fitted models, their standard errors (SEs), and the previous measures are given in Tables 6 and 7 for both data sets. The numbers in these tables indicate that the MOWE distribution gives a superior fit over the other models tested. The PDF, CDF, survival function (SF) and probability–probability (PP) plot for the MOWE model are reported in Fig. 2 for both data sets. Fig. 3 provides the total time on test (TTT) plots for both data sets and it also illustrates that the HRF of the first data is increasing because it has a concave shape. The HRF of the second data is decreasing because the TTT plot has a convex shape. Hence,

the MOWE distribution can capture all data sets with monotone HRF properly.

7. Concluding remarks

We constructed a new competitive family of distributions to the well-established beta-G and Kumaraswamy-G classes. Some of its mathematical properties were determined. We addressed eight estimation methods for a special model called the MOW-exponential (MOWE) distribution. The simulation results showed that the maximum likelihood approach is the best estimation method for the MOWE parameters. We proved the utility of this distribution to analyze COVID-19 data from Canada and China.

The topics of this article can be extended in several ways. For example, a discrete version of the new family can be established and its properties can be explored. Bivariate extensions of the new family can also be investigated.

Data Availability

This work is mainly a methodological development and has been applied on secondary data, but, if required, data will be provided.

Fund

This study was funded by Taif University Researchers Supporting Project number (TURSP-2020/279), Taif University, Taif, Saudi Arabia.

Declaration of Competing Interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

References

- Afify, A.Z., Al-Mofleh, H., Dey, S., 2021a. Topp-Leone odd log-logistic exponential distribution: Its improved estimators and applications. *Anais da Academia Brasileira de Ciências* 93, e20190586.
- Afify, A.Z., Cordeiro, G.M., Ibrahim, N.A., Jamal, F., Elgarhy, M., Nasir, M.A., 2021b. The Marshall-Olkin odd Burr III-G family: theory, estimation, and engineering applications. *IEEE Access* 9, 4376–4387.
- Afify, A.Z., Marzouk, W., Al-Mofleh, H., Ahmed, A.H.N., Abdel-Fatah, N.A., 2022. The extended failure rate family: properties and applications in the engineering and insurance fields. *Pakis. J. Stat.* 38(1), 165–196.
- Afify, A.Z., Yousof, H.M., Alizadeh, M., Ghosh, I., Ray, S., Ozel, G., 2020. The Marshall-Olkin transmuted-G family of distributions. *Stochast. Qual. Control* 35, 79–96.
- Al-Babtain, A.A., Shakhathreh, M.K., Nassar, M., Afify, A.Z., 2020. A new modified Kies family: properties, estimation under complete and type-II censored samples, and engineering applications. *Mathematics* 8, 1345.
- Alexander, C., Cordeiro, G.M., Ortega, E.M., Sarabia, J.M., 2012. Generalized beta-generated distributions. *Comput. Stat. Data Anal.* 56, 1880–1897.
- Bourguignon, M., Silva, R.B., Cordeiro, G.M., 2014. The Weibull-G family of probability distributions. *J. Data Sci.* 12, 53–68.
- Chen, Z., 2000. A new two-parameter lifetime distribution with bathtub shape or increasing failure rate function. *Stat. Prob. Lett.* 49, 155–161.
- Cordeiro, G.M., Afify, A.Z., Ortega, E.M., Suzuki, A.K., Mead, M.E., 2019. The odd Lomax generator of distributions: properties, estimation and applications. *J. Comput. Appl. Math.* 347, 222–237.

- Cordeiro, G.M., de Castro, M., 2011. A new family of generalized distributions. *J. Stat. Comput. Simul.* 81, 883–898.
- Cordeiro, G.M., Lemonte, A.J., Ortega, E.M., 2014. The Marshall-Olkin family of distributions: mathematical properties and new models. *J. Stat. Theory Practice* 8, 343–366.
- Cordeiro, G.M., Ortega, E.M., da Cunha, D.C., 2013. The exponentiated generalized class of distributions. *J. Data Sci.* 11, 1–27.
- Eugene, N., Lee, C., Famoye, F., 2002. Beta-normal distribution and its applications. *Commun. Stat. Theory Methods* 31, 497–512.
- Gupta, R.C., Gupta, P.L., Gupta, R.D., 1998. Modeling failure time data by Lehman alternatives. *Commun. Stat. Theory Methods* 27, 887–904.
- Hassan, A.S., Hameda, S.E., 2017. A new family of additive Weibull-generated distributions. *Int. J. Math. Appl.* 4, 151–164.
- Jones, M.C., 2004. Families of distributions arising from distributions of order statistics. *Test* 13, 1–43.
- Khan, M.S., King, R., Hudson, I.L., 2017. Transmuted generalized exponential distribution: a generalization of the exponential distribution with applications to survival data. *Commun. Stat. Simul. Comput.* 6, 4377–4398.
- Korkmaz, M.C., Cordeiro, G.M., Yousof, H.M., Pescim, R.R., Afify, A.Z., Nadarajah, S., 2019. The Weibull Marshall-Olkin family: regression model and application to censored data. *Commun. Stat. Theory Methods* 8, 4171–4194.
- Lemonte, A.J., Cordeiro, G.M., Moreno-Arenas, G., 2016. A new useful three-parameter extension of the exponential distribution. *Statistics* 50, 312–337.
- Liu, X., Ahmad, Z., Gemeay, A.M., Abdulrahman, A.T., Hafez, E.H., Khalil, N., 2021. Modeling the survival times of the COVID-19 patients with a new statistical model: a case study from China. *PLoS One* 16, e0254999.
- Marshall, A.W., Olkin, I., 1997. A new method for adding a parameter to a family of distributions with application to the exponential and Weibull families. *Biometrika* 84, 641–652.
- Mead, M.E., Afify, A.Z., Butt, N.S., 2020. The modified Kumaraswamy Weibull Distribution: properties and Applications in Reliability and Engineering Sciences. *Pakistan J. Stat. Oper. Res.* 16, 433–446.
- Nadarajah, S., Kotz, S., 2006. The exponentiated type distributions. *Acta Appl. Math.* 92, 97–111.
- Nassar, M., Afify, A.Z., Dey, S., Kumar, D., 2018. A new extension of Weibull distribution: properties and different methods of estimation. *J. Comput. Appl. Math.* 336, 439–457.
- Nassar, M., Kumar, D., Dey, S., Cordeiro, G.M., Afify, A.Z., 2019. The Marshall-Olkin alpha power family of distributions with applications. *J. Comput. Appl. Math.* 351, 41–53.
- Nofal, Z.M., Afify, A.Z., Yousof, H.M., Cordeiro, G.M., 2017. The generalized transmuted-G family of distributions. *Commun. Stat. Theory Methods* 46, 4119–4136.
- Phani, K.K., 1987. A new modified Weibull distribution. *Commun. Am. Ceram. Soc.* 70, 182–184.
- Ramos, P.L., Nascimento, D.C., Cocolo, C., Nicola, M.J., Alonso, C., Ribeiro, L.G., Ennes, A., Louzada, F., 2018. Reliability-centered maintenance: analyzing failure in harvest sugarcane machine using some generalizations of the Weibull distribution. In: *Modelling and Simulation in Engineering*, 2018, p. 1241856.
- Ramos, P.L., Nascimento, D.C., Ferreira, P.H., Weber, K.T., Santos, T.E., Louzada, F., 2019. Modeling traumatic brain injury lifetime data: improved estimators for the generalized gamma distribution under small samples. *PLoS one* 14, e0221332.
- Rasekhi, M., Alizadeh, M., Altun, E., Hamedani, G.G., Afify, A.Z., Ahmad, M., 2017. The modified exponential distribution with applications. *Pakistan J. Stat.* 33, 383–398.
- Ristic, M.M., Kundu, D., 2015. Marshall-Olkin generalized exponential distribution. *Metron* 73, 317–333.
- Rodrigues, G.C., Louzada, F., Ramos, P.L., 2018. Poisson-exponential distribution: different methods of estimation. *Journal of Applied Statistics* 45, 128–144.
- Shaw, W.T., Buckley, I.R., 2007. The alchemy of probability distributions: beyond Gram-Charlier expansions, and a skew-kurtotic-normal distribution from a rank transmutation map, UCL discovery repository. <http://discovery.ucl.ac.uk/id/eprint/643923>.
- Smith, R.M., Bain, L.J., 1975. An exponential power life testing distribution. *Commun. Stat.—Theory Methods* 4, 469–481.
- Tahir, M.H., Nadarajah, S., 2015. Parameter induction in continuous univariate distributions: well-established G families. *Anais da Academia Brasileira de Ciências* 87, 539–568.
- Yousof, H.M., Afify, A.Z., Hamedani, G.G., Aryal, G.R., 2017. The Burr X Generator of Distributions for Lifetime Data. *J. Stat. Theory Appl.* 16, 288–305.
- Zaidi, S.M., Sobhi, M.M.A., El-Morshedy, M., Afify, A.Z., 2021. A new generalized family of distributions: properties and applications. *Aims Math.* 6, 456–476.
- Zografos, K., Balakrishnan, N., 2009. On families of beta-and generalized gamma-generated distributions and associated inference. *Stat. Methodol.* 6, 344–362.